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Sustained Effect on Human Capital?**

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# Can Early Intervention have a Sustained Effect on Human Capital? \*

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## ABSTRACT

*Evidence on the sustained effect of early intervention is inconclusive, with many studies experiencing a dissolution of treatment effects once the program ends. Using a randomized trial, this paper examines the impact of Preparing for Life (PFL), a pregnancy to age five home visiting and parenting program, on outcomes in middle childhood. We find little evidence of cognitive fade-out at age nine, with significant treatment effects on cognitive skills (0.67SD) and school achievement tests (0.47-0.74SD) that are of a similar magnitude to those observed at the end of the program. There is no impact on other school outcomes and earlier effects for socio-emotional skills are no longer evident. While about 50 percent of the sample is retained at age nine, the treatment groups are still balanced on all key baseline characteristics and the results are robust to inverse probability weighting. Mediation analysis suggests that ~46 percent of the treatment effect on cognitive skills is explained by improvements in early parental investment. This study demonstrates that boosting children's early cognitive skills can reduce school-age inequalities five years after program completion, yet continued investment may be needed to break long-standing inequalities in other dimensions of skills.*

**Keywords:** Early childhood intervention; cognitive skills; socio-emotional and behavioral skills; randomized control trial; school-age inequalities.

**JEL Classification:** C93, D13, I26, J13

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## I Introduction

There has been a sizable increase in the number of studies in economics focusing on the early childhood period. This is in recognition of the importance of this period for human capital formation. Biologically, pregnancy and the earliest years are the most productive period for brain development, whereby infants' underlying neurodevelopmental capacities are formed (Cantor *et al.* 2019; Knudsen *et al.* 2006; Thompson and Nelson 2001). Although this process is largely genetic, research from epigenetics demonstrates that the environment, and in particular the level of parental stimulation and sensitive caregiving, can shape the strength of structural and functional brain networks in either a positive or negative way (Shonkoff 2010). This is important as these neurological processes are instrumental in shaping children's later cognitive, social, behavioral, and physical development (Blair and Raver 2012).

Inequalities in children's capacities typically arise in contexts of disadvantage, where a family's ability to invest in their children may be hampered by monetary or cognitive constraints (Becker 1965; Mani *et al.* 2013) or the stress that accompanies poverty (Conger *et al.* 1994; Lupien *et al.* 2001; Masarik and Conger 2017). These inequalities develop early in life and unless remediated continue to widen as children progress through school (Cunha and Heckman 2007; Heckman and Mosso 2014). Duncan *et al.* (2019) notes that rising income inequality in the US in recent years has been accompanied by a widening gap in achievement scores across social groups, and Waldfogel and Washbrook (2011) find that between a third and a half of the income-related gap in children's cognitive ability may be attributed to differences in parenting style and the home learning environment. Yet despite growing evidence on the importance of parental investment in the prenatal and infancy period in particular (Almond and Currie 2011), there remains a dearth of studies identifying effective interventions to circumvent these later inequalities.

This paper aims to address this gap by examining the impact of an early intervention program at reducing school-age inequalities in children's skills. The trial, known as *Preparing for Life (PFL)*, began in 2008, and is based on the premise that investing in intensive parenting supports from pregnancy until age five will permanently alter children's skill level. The study targeted pregnant women residing in a highly disadvantaged suburban community in Ireland and randomly allocated them to a 'high' or 'low' treatment group. The high treatment group received fortnightly home visits from trained mentors from pregnancy until age five, in addition to baby massage classes in the first year and group-based parenting classes in the second year. The cohort were assessed at eight time points during the trial and by age five, Doyle (2020) finds that the program significantly improved

children's cognitive, social, and behavioral development.<sup>1</sup> Specifically, children in the high treatment groups had higher language ability (0.67 of a standard deviation, SD), spatial ability (0.65 SD), and pictorial reasoning ability (0.56 SD) compared to children in the low treatment group. They were also significantly less likely to score below average (20 vs 60 percent) and significantly more likely to score above average (25 vs 8 percent) in terms of overall cognitive ability. The program had some impact on socio-emotional skills, although the effect sizes were smaller and the results less robust. For example, the high treatment children were less likely to experience clinically significant externalizing (0 vs 16 percent) and internalizing (3 vs 20 percent) problems at age four, and engaged in better prosocial behavior (Doyle 2020), however there were no effects on teacher-reported scores of socio-emotional skills in the first year of primary school (Doyle and *PFL* Evaluation Team 2016).

The contributions of this paper are two-fold. First, to test for the continuity or fade-out of early intervention treatment effects on human capital in middle childhood. The objective of the *PFL* program was to improve children's school readiness skills with the expectation that boosting children's skills prior to school entry would allow them to fully exploit the learning opportunities which school provides. Thus, by assessing the impact of the program at age nine, we can document its success or otherwise at narrowing socio-economic gaps in school-age skills. Data are available on cognitive ability, which is considered a relatively stable trait, and achievement tests which are more amenable to schooling and other investments (Magnuson and Duncan 2018), as well as different dimensions of non-cognitive skills capturing both cognitive and behavioral components. Analysing multiple measures of skills allows us to identify the areas through which early intervention can have a persistent effect.

The evidence base used to justify investment in the early years is founded on a handful of experiments which were conducted in the 1960's and 70's and continue to follow participants into adulthood. The results from these landmark studies, such as the Perry Preschool, Abecedarian, and Jamaica studies, typically find that intervening early in life leads to higher earnings and employment, reduced crime, and better health in adulthood (e.g. Campbell and Ramey 1994; Gertler *et al.* 2014; Heckman *et al.* 2010; Heckman *et al.* 2013; Schweinhart 2013). Many of the more recent interventions, particularly the new wave of home visiting programs operating in developing countries

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<sup>1</sup> A number of papers published during the trial report interim findings. Doyle (2013) sets out the design of the study. Doyle *et al.* (2014) analyze maternity hospital data and finds that the program had a significant effect on reducing caesarean sections. Doyle *et al.* (2017a) finds that the program had no impact on early cognitive and non-cognitive skills measured at six, 12, and 18 months, although effects on some parenting outcomes and measures of the home environment were significant. Doyle *et al.* (2015) focused on child health at five time-points between six and 36 months find effects on respiratory illnesses and general health problems, but only at 24 months. O'Sullivan, Fitzpatrick and Doyle (2017) conduct a mediation analysis and finds that improvements in dietary intake, and protein in particular, mediates the impact of the program on cognitive development at 24 and 36 months. Doyle *et al.* (2017b) find no evidence that the program improved maternal well-being. Cote *et al.* (2018) investigate whether the impact of the program varies according to children's developmental trajectories and find a positive impact on trajectories of cognitive development and number of health clinic visits for all children, whereas positive impacts on externalizing behavior problems are restricted to children with the most severe problems. Doyle (2020) reports on the final outcomes of the trial.

(usual tied to conditional cash transfer programs), either do not follow participants beyond the intervention period or will have to wait several years to identify school age effects (e.g. Andrew *et al.* 2018; Attanasio *et al.* 2020; Barry *et al.* 2017).

This has resulted in what Almond, Currie and Duque (2018) refer to as the ‘missing middle’. They argue that the full impact of early shocks, and by extension, early investment, in childhood is not observed until adulthood, and there is a relative lack of knowledge on the mechanisms through which policies implemented in early childhood transition through middle childhood and into adulthood. For example, many of the landmark studies experienced a fading out of cognitive effects in the aftermath of the intervention, yet improved outcomes in adulthood (e.g. Campbell *et al.* 2014; Heckman *et al.* 2017). This points to either a latency/sleeper effect or that the intervention impacts other unmeasured skills in childhood and adolescence which are important for later development (Almond *et al.* 2018). Work by Heckman and colleagues, using the Perry Preschool program, points to the latter and finds that treatment effects on earnings, criminality and health are largely driven by improvements in non-cognitive skills, in particular externalizing behavior (Heckman, Pinto, and Salvelyev 2013).

To date, there are few follow-up studies of more contemporary interventions which focus on the prenatal and infancy period. While the seminal studies are influential in shaping our understanding of the long term impact of intervening in an environment that was largely devoid of other family supports, the ‘services as usual’ or counterfactual for children participating in contemporary trials is typically more generous in terms of subsidized childcare and financial supports for low-income families, at least in developed countries. Thus, on the one hand, we would anticipate lower effect sizes in contemporary trials, both during and following the intervention, as high risk families receive more ‘treatment’ as standard practice. On the other hand, effect sizes may be higher and more persistent in contemporary trials if dynamic complementarities exist (Cunha and Heckman 2007) and treatment children, whose early skills have been boosted by the intervention during infancy can exploit the higher level of universal investments made in school-age children. This study can help to address this hypothesis by investigating whether the positive effects observed at the end of the *PFL* trial persist or dissipate in later childhood and whether the effects are stronger for different types of skills. In this way this study starts to fill the ‘missing middle’ by examining the impact of intervening during the earlier stages of life on human capital at nine years old.

The second contribution of this paper is to identify the mechanisms through which treatment effects operate. Most early intervention programs are based on the premise that providing children with a supportive environment helps to protect them against the risk factors (such as poverty and living in a high stress environment) which can compromise healthy development (Shonkoff 2010). While interventions differ in their approach (e.g. center-based programs, home visiting programs, conditional/unconditional cash transfer programs) they all promote caregiving behavior that is based

on the provision of enriching and stimulating environments which is responsive to children's developmental needs. The intervention studied in this paper is primarily a home visiting program, where the specific focus is on improving parenting behavior and the quality of the home environment. Home visiting programs operate by changing the knowledge, attitudes, and beliefs of parents, as well as encouraging greater parent-child interactions (Britto *et al.* 2015). Thus the mechanism through which the intervention should impact child outcomes is via changes in parental behavior. As *PFL* is a holistic intervention with multiple treatments and whose remit within the broad scope of 'parenting behavior' is wide, identifying the mechanisms through which treatment effects emerge is complicated by the sheer number of factors upon which the intervention could impact. We address this using exploratory factor analysis to summarize 190 different measures of parenting behaviors collected between the ages of six and 48 months. We then use the resultant factors to conduct a mediation analysis to determine what proportion, if any, of the age nine treatment effects can be explained by improvements in early parenting behavior.

A common concern in longitudinal trials is attrition. In this study, approximately 65 percent of the sample participated in the age five assessment, with equal representation across the high and low treatment groups. By age nine, we retained ~60 percent of the high treatment group and ~40 percent of the low treatment group. Despite this unbalance in the number of participants in the groups, tests for baseline equivalence shows that the treatment groups are still balanced on all key baseline characteristics, indicating that the randomization assumption still holds. In addition, all results are estimated both with and without inverse probability weights and are robust to their inclusion.

Overall, we find little evidence of cognitive fade-out at age nine, with effect sizes of 0.67 SD on general conceptual ability and standardized school achievement tests of reading (0.74 SD) and math (0.47 SD). These effects are larger than those found in both the seminal studies of home visiting programs (e.g. Heckman *et al.* 2017) and the early outcomes that have emerged from more contemporary home visiting trials (e.g. Attanasio *et al.* 2014). The program, however, has no impact on absenteeism or the use of school resources and the significant treatment effects observed for children's socio-emotional skills and behavior at age four are no longer present at age nine. Mediation analysis suggests that ~46 percent of the treatment effect on cognitive skills is explained by improvements in early parenting behavior. The large and persistent effect on the more cognitive dimensions of children's skills may be attributed to both the length and intensity of the *PFL* program. Many of the seminal and more recent interventions are typically shorter in duration and start later in the lifecycle. By intervening in pregnancy and infancy, *PFL* helps to optimize brain development during a period of heightened malleability, and by continuing the investment until school age it exploits this elevated skill level to develop more advanced skills. The lack of persistent effects on the non-cognitive dimensions of prosociality and behavior may be attributed to the weaker treatment effects and the concentration of effects within the clinical range found at earlier ages.

The remainder of the paper is organized as follows. Section II sets out the theoretical background and reviews the literature on school-age follow-ups of early intervention programs. Section III describes the intervention and study procedures. Section IV presents the results. Section V summarizes and concludes.

## **II Background & Literature**

The Cunha and Heckman (2013) model of skill formation can be used to hypothesize the likely impact of early intervention programs in middle childhood. Their model captures the role of child endowments and parental/caregiver investment in shaping human capital over time. It posits that mastering simple skills is a prerequisite for mastering more complex skills, a process referred to as self-productivity. For example, a child must recognize numbers before performing addition. Thus, if an intervention is effective in raising children's basic skills early in life, this facilitates the acquisition of more advanced skills later in life. Bailey *et al.* (2017) argue that interventions which develop skills incrementally and target fundamental skills which are predictive of later skills have a higher probability of achieving persistent effects. Based on this theory, one would expect that providing sustained investment throughout the critical periods of skill formation which can respond to the child's growing skill set will be more likely to demonstrate effects beyond the life of the intervention. Thus, the *PFL* program, which operates from pregnancy until age five and is based on a model of supporting and coaching parents to optimize their child's development using a set of age-appropriate tip sheets, may be more likely to generate lasting change than center-based programs which start later in childhood thus missing the critical 'first 1000 days', and due to higher child-caregiver ratios, may not have the flexibility to respond to children's individual needs.

Cunha and Heckman (2013) also state that the skills developed in one period make investments in subsequent periods more effective, a process known as dynamic complementarity. Essentially, if an intervention raises skills early in life, this allows children to capitalize on later investments. For example, children who start school with greater word recognition and vocabulary skills are more likely to benefit from school-based literacy instruction compared to children lacking these basic skills. This later investment may come from targeted interventions (e.g. preschool support for disadvantaged children) or universal services (e.g. public primary schools). Bailey *et al.* (2017) note that while the quality of this later investment may impact children's skills, the gap between the treatment and control groups will remain or widen if dynamic complementarities are at play. In the context of *PFL*, as described below, there are no differences in the type or quality of schools attended by the treatment and control groups, thus if the treatment group exhibits superior skills at age nine, it would provide some support for this skill beget skill hypothesis.

Despite these theoretical considerations, much of the evidence base for early intervention programs tends to exhibit a pattern of fade-out rather than persistence over time. A study by Baily *et*

*al.* (2017) examines the persistence of 67 high-quality early intervention programs<sup>2</sup> and finds a general pattern of declining effects sizes, with end of program effects averaging 0.23 SD, which then fall to 0.10 SD by the end year one, and 0.05 SD up to two years after treatment has ended.<sup>3</sup> They note that apart from the few seminal programs, the majority of studies fail to follow participants thus precluding the analysis of long term effects.

Regarding the seminal studies, most of which are center-based programs starting later in childhood, the majority experience a fade-out of their initial treatment effects. For example, the Perry Preschool Program found that the large and significant impact on cognitive development of 0.75 SD identified at the end of the program had fallen to 0.08 SD by age eight (Schweinhart *et al.* 2005). In addition, the effect sizes of ~0.20 SD on cognitive skills identified at the end of the Head Start program (age four) were no longer present at subsequent follow-ups between kindergarten and third grade (eight-nine years old) (Puma *et al.* 2012). Despite such cognitive fade-outs, these programs still observed improvements in other outcomes later in life such as involvement in crime and being in receipt of social welfare (Heckman *et al.* 2017). In addition, some did not report such cognitive fade-outs. For example, the Jamaica study, which was based on weekly home visits from nine to 24 months, found that treatment effects on IQ of 0.88 SD had dissipated by age seven, but reemerged at the 11, 17, and 22 year follow-ups with effect sizes ranging between 0.40 - 0.60 SDs (Grantham-McGregor and Smith 2016). In addition, the significant IQ effect of 0.74 SD in the Abecedarian program, which offered center-based care and home visits from infancy until age five, persisted, although declined to 0.37 SD, on average, during the follow-up periods at ages eight, 12, 15, and 21 (Campbell *et al.* 2001). It is notable that the studies which experienced early cognitive fade-out are based on programs starting at age three (e.g. Perry and Head Start), while programs starting earlier in the lifecycle appear to demonstrate more persistent effects (e.g. Jamaica and Abecedarian), although this observation is based on a handful of studies with long-term follow-ups.

Regarding contemporary interventions specifically focusing on home-based programs starting before or shortly after birth, the evidence base for sustained effects is also inconsistent. Systematic reviews reporting on treatment effects during or shortly after the program's end, find impacts on children's health, development, behavioral problems, family economic self-sufficiency, and positive parenting practices (Filene *et al.* 2013; Peacock *et al.* 2013; Sweet and Appelbaum 2004), however these effects tend to be short lived (Burger 2010; Peacock *et al.* 2013) and few continue to follow the cohorts into middle childhood. Appendix Tables 1a-c summarize the sparse literature examining the

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<sup>2</sup> Note that the majority of these studies are based on center-based early childhood education programs, with one-third being assessed through random assignment, and the remaining using some form of quasi-experimental design.

<sup>3</sup> A meta-analysis focusing on programs targetting early phonological awareness, conducted by Bus and van IJzendoorn (1999), found large initial impacts on children's reading skills (0.44 SD) which faded to 0.16 SD at the 18 month assessment on average. A number of mathematic interventions also find that large initial gains tend to fade over time (Clements *et al.* 2013; Smith *et al.* 2013).



medium term impact of home visiting programs on children's outcomes between the ages of five and 12 in developed countries, thus broadly corresponding to the follow-up period considered in this paper.

Table 1a shows that of the four studies assessing the medium-term impact on children's cognitive development, only one has a significant effect. In particular, Bierman *et al.* (2017) find that children who participated in Early Head Start (home based component) had improved cognitive ability at ages seven to nine. However, no effects are found for Early Head Start at age five (Chazan-Cohen, Raikes and Vogel 2013) or the Nurse Family Partnership program at ages six-nine (Olds *et al.* 2014) or age 12 (Kitzman *et al.* 2010).

Table 1b shows that of the five studies assessing children's achievement tests, three identify a significant treatment effect. Bierman *et al.* (2017) also find that Early Head Start had an impact on children's reading and language skills at ages seven to nine, and two studies of the Healthy Families America program identify significant effects on the percentage of children in a gifted program, receiving special education, and excelling academically in behaviors that promote learning at ages six to seven (DuMont *et al.*, 2010; Kirkland and Mitchell-Herzfield 2012). The Nurse Family Partnership program had no impact on achievement tests at ages six to 12 (Sidora-Arcoleo *et al.* 2010) or at age 12 (Kitzman *et al.* 2010).

Finally, Table 1c shows that of the ten studies assessing the medium-term impact of home visiting programs on children's socio-emotional skills and behavior, only three identify significant effects. In particular, two studies of the Early Head Start program find effects on children's behaviors, perceived competence, and approaches to learning at ages five (Chazan-Cohen, Raikes, and Vogel 2013) and ages seven to nine (Bierman *et al.* 2017). In addition, the Nurse Family Partnership program finds a reduction in internalizing disorders at age 12 (Kitzman *et al.* 2010). However, a number of other home visiting studies, including Healthy Families America at age seven (DuMont *et al.* 2010), Healthy Steps at age five (Minkowitz *et al.* 2007), Family Check-Up at age five (Sitnick *et al.* 2015) and age seven to eight (Smith *et al.* 2014), as well as the Nurse Family Partnership program at ages six to nine (Olds *et al.* 2014) and ages six to 12 (Sidora-Arcoleo *et al.* 2010), fail to identify sustained effects on children's socio-emotional skills and behavior.

In general, based on this literature, there is little evidence that home visiting programs have a sustained impact on children's development in middle childhood. It is important to note that many of these programs demonstrated small, although significant, effects at the end of the program, thus it is possible that interventions which achieve sizeable improvements across multiple dimensions of skills at program completion have a greater chance of demonstrating longer term impacts.

### III Methods

#### A. The *PFL* Intervention

The intervention considered in this paper is based on the *PFL* program. The program was developed in a bottom-up fashion by local service providers and community representatives with the aim of advocating for greater parental investment in children. The over-arching aim of the program was to reduce social inequalities in children's skills by working with families residing in a highly disadvantaged community from pregnancy until approximately age five. The program was delivered by a team of family mentors who were directly employed on a full-time basis. Although the mentors came from different professional backgrounds, ranging from social work to psychology, they all received six months training before the trial began, as well as monthly supervision throughout the trial, to ensure consistency of program delivery.

*PFL* was essentially a dosage experiment such that all participants received some low levels supports and the 'high' treatment group received additional intensive parenting supports. The 'low' treatment included five developmental-toy packs containing, for example, building blocks and a play mat; four developmental-book packs containing between six-eight books; two framed professional photographs of the target child; access to a support worker to assist in non-parenting related issues; invitations to attend workshops on healthy eating and stress control, as well as *PFL* specific social event (e.g. Christmas parties, coffee mornings). The 'high' treatment included a baby massage course, the Triple P Positive Parenting Program, and the primary treatment - a five year home visiting program. Each treatment was designed to support and encourage parents to engage in parenting behaviors which would optimize their children's development.

Although the evidence base for baby massage is limited (see Bennett, Underdown, and Barlow 2013), participants in the high treatment group were invited to attend five two-hour baby massage classes during the first year of the program. The dual aim of these sessions were to encourage early physical and gaze interactions between parents and their infants, as well as serve as an incentive for early engagement with the *PFL* program. In total, 62 percent of the high treatment group attended at least one of the classes.

All the *PFL* mentors were trained in delivery of the *Triple P Positive Parenting Program*. Triple P is a universal program which operates at five levels of intervention including a media campaign and communication strategy, a positive parenting seminar series, single session discussion groups, intensive small group and individual programs, and intensive family intervention (Sanders, Markie-Dadds, and Turner 2003). Its core principle is the promotion of positive parenting strategies through setting clear rules and boundaries while providing a safe and engaging home environment (Sanders 2012). Each Triple P level has been subject to some form of experimental evaluation, with demonstrated evidence of effect on parenting behavior and children's socio-emotional skills (Sanders

*et al.* 2014) The *PFL* mentors used the Triple P strategies throughout the home visits and during the third year of treatment (between 24 and 36 months) participants were invited to take part in five group-based Triple P sessions. These two-hour sessions took place in the local community center and were facilitated by the mentors who worked through a range of goal setting exercises focused on managing misbehaviour and promoting child development using active discussion, videos, vignettes and role playing. In total, 43 percent of the high treatment group took part in all least one Triple P session.

At its core *PFL* is a home visiting program. Each family in the high treatment group was assigned a *PFL* mentor to work with the family in their home from pregnancy until their child started primary school. It was hypothesized that working with the same family for such an extended and critical period in the child's life would facilitate a relationship of trust and respect to develop, which in turn, would ensure the efficacious delivery of the program's content. Although the prescribed treatment was based on one-hourly home visit every two weeks, there was considerable variability in the quantity of visits delivered with 16 percent of families participating in more than 90 home visits and 17 percent participating in no visits. On average, families received 50 visits from the mentors (SD = 38. range 0 - 145), which represents approximately one visit per month.

The mentoring was based on three core principles of providing knowledge and active guidance on appropriate parenting techniques, helping parents to identify and promote children's developmental milestones, and encouraging parents to provide greater stimulation to their children, particularly during infancy. The mentors used an active learning approach which involved modelling behavior which promotes child development, observation and feedback on parents' interactions with their children, as well as providing skills training in various domains such as nutrition and discipline. The success of these strategies was dependent on the strength of the mentor-parent relationship, thus the mentors invested a significant amount of resources, especially during the first year, on building this relationship.

The content of the home visits was based ~200 'Tip Sheets' which were delivered based on the age of the child and the needs of the family. The Tip Sheets delivered between pregnancy and age one focused on pregnancy health, preparing for the birth, nutrition and infant health, providing a safe home environment, building a secure parent-infant relationship, and methods for promoting the infant's early cognitive development, amongst others. The Tip Sheets delivered between age one and two continued to focus on child health and nutrition, but with a larger emphasis on supporting the child's socio-emotional development, as well as identified and promoting developmental milestones. The final set of Tip Sheets delivered between age two and school entry, continued to focus on methods for developing the child's cognitive and socio-emotional development, with a particular emphasis on language skills and managing challenging behavior, as well as preparing the child for school. While the Tip Sheets were developed by the *PFL* program team, they were informed by best practice evidence on child development and parenting. The mentors used a variety of tools to deliver

the material including role modelling and observation techniques, discussion and feedback, and critical review and goal setting. A qualitative study of the *PFL* program identified the importance of the parent-mentor relationship and program flexibility as the main factors influencing participant engagement with the program (Lovett *et al.* 2016).

## **B. Study design**

The *PFL* program was delivered between 2008 and 2015 to pregnant women living a high-poverty community in Dublin, Ireland.<sup>4</sup> Apart from geographical residency, there were no exclusion criteria. Women were recruited through the local maternity hospital and a community-based advertising campaign. As recruitment was voluntary, an estimated 52 percent of all eligible women were recruited into the study.<sup>5</sup> After consenting to join the study, participants were randomly assigned to either the high treatment group (n=115) or the low treatment group (n=118) using individual-level randomization with no blocking or stratification.<sup>6</sup> Randomization was conducted on a tablet computer whereby participants were invited to initiate their own treatment assignment by tapping the *PFL* logo on the welcome page. This action initiated an email listing the participant's ID number and treatment condition which was automatically sent to the study team and acted as a permanent record of treatment assignment, thus countering any subversion by the recruitment officer or the participant. A baseline comparison of the high and low treatment groups found that there were no statistically significant differences between the groups on 108 of the 117 measures tested, confirming the success of the randomization procedure. Please refer to Doyle (2013) and Doyle (2020) for more detail on the recruitment and randomization procedures.

Outcome data were collected from participants at multiple time points during the trial (pregnancy, and when the children were six, 12, 18, 24, 36, 48, 51, and 60 months).<sup>7</sup> Five years after the program ended, additional funding was received to conduct a follow-up of the cohort at age nine. All families who were originally recruited and randomized were eligible to take part in the follow-up however, participants who had officially dropped out or left the study due to death or miscarriage

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<sup>4</sup> The community is characterized by high rates of unemployment (12 percent) and low education (7 percent college degree), as well as high rates of public housing (42 percent) (Census 2006) and poor levels of children's school readiness skills (Doyle *et al.* 2012).

<sup>5</sup> Among those who did not join the study, the study team had no contact with approximately half of eligible non-participants and had some level of contact with the other half who either refused to join the study or could not be further contacted for recruitment. A survey of eligible non-participants, conducted through the local preschool centers when the children were four years old, found there were no differences between participants and non-participants in terms of maternal age, family size, parity, relationship status, or type of employment, however participants were younger at the birth of their first child, had lower levels of education, were less likely to be employed, and were more likely to be eligible for free medical care compared to non-participants. Please see Doyle (2020) for more information.

<sup>6</sup> A power analysis, based on 80 percent power, a 5 percent significance level, and a minimum detectable effect of 0.18 SD for cognitive skills (based on a meta-analysis of home visiting programs by Sweet and Appelbaum 2004), found that a sample of 117 in the treatment and control groups was required.

<sup>7</sup> All *PFL* data is publicly available for use and can be accessed in the Irish Social Data Archive at [www.ucd.ie/issda](http://www.ucd.ie/issda)

were not contacted. In the first phase of recruitment, all families were invited to attend a community event organised by the *PFL* mentoring team and eligible attendees were invited to participate in the follow-up. As a second phase of recruitment, eligible participants who did not attend the community event and who had previously agreed to future contact by the study team, were contacted directly and invited to participate. In both phases of recruitment, eligible participants were provided with information about the follow-up study and were asked to re-consent to join the study.

### **C. Attrition**

Figure 1 depicts the families' participation in the study between program entry and the age nine follow-up. At the end of the program, 65 percent of the high treatment group and 63 percent of the low treatment group were retained, thus attrition was largely equivalent across both groups. However, by age nine, attrition was significantly higher among the low treatment group. Child assessment data were collected for 117 of the original 233 randomly assigned participants, representing an overall retention rate of 50 percent (high=59 percent; low=41 percent;  $p$ -value=0.003). For the school data, information on 123 participants were available, representing a retention rate of 53 percent (high=61 percent; low=45 percent;  $p$ -value=0.015). For the parent data, surveys were conducted by 111 participants, representing an overall retention rate of 48 percent (high=56 percent; low=40 percent;  $p$ -value=0.016). This higher rate of attrition among the low treatment group may, in part, be attributed to the recruitment strategy which largely depended on the *PFL* mentors, with whom the high treatment families have greater contact. As shown in Appendix Table 1a-c, these rates are somewhat lower than the follow-up studies of the mainly US-based home visiting programs.<sup>8</sup>

A comparison of the high and low treatment groups at baseline using the age nine estimation samples finds that the groups significantly differ (at the 10 percent level) on 6.8 percent (8/117) of baseline measures for the child assessment data, 10.3 percent (12/117) of measures for the school data, and 12.8 percent (15/117) of measures for the parent data. These are largely consistent with chance and indicate that the groups remain balanced at the follow-up, particularly for the child data.<sup>9</sup> Table 1 compares the age nine participants in the high and low treatment groups across the three estimation samples for a selection of baseline variables. It shows that there are no statistically significant differences across the two groups for a range of socio-demographic factors including parental age, education, employment status, and health.

Although the estimation samples are largely balanced in terms of baseline characteristics, it is important to test for differential attrition in the high and low treatment group. To investigate this, the

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<sup>8</sup> The introduction of the General Data Protection Regulation (GDPR) in 2018, an EU law on data protection and privacy, limited our ability to directly contact the original trial participants without explicit consent beyond the life of the trial.

<sup>9</sup> As discussed in results section, the results are robust to conditioning on baseline differences.

factors predicting participation in the child, school, and parent assessments are tested separately for the high and low treatment groups using bivariate tests with 50 of the baseline measures. We find some evidence of differential attrition, with between 17 and 36 percent of baseline measures predicting attrition from the high treatment group at age nine, and between 11 and 19 percent of measures predicting attrition from the low treatment group (in two-tailed tests, using the 10 percent significance level).<sup>10</sup> Overall, there is less evidence of differential attrition in the school sample than the parent or child samples. The factors associated with attrition are similar in the high and low treatment groups, however the number of variables predicting attrition is lower in the low treatment group. Table 2 compares a selection of baseline characteristics of those who participated in the age nine parent assessment, i.e. stayers, to those who did not, i.e. non-stayers. It shows that high treatment parents who completed the age 9 assessment were older, had higher IQ, and were more likely to be employed at baseline, while low treatment parents who completed the age nine assessment were less likely to be first time mothers.

In order to account for differential attrition across the high and low treatment groups, the treatment effects are estimated using the Inverse Probability Weighting (IPW) procedure (Robins, Rotnitzky, and Zhao 1994). To determine which baseline measures predict the age nine drop-out rate, separate bivariate tests are conducted and the Bayesian Information Criterion (BIC; Schwarz 1978) is used to reduce down the number of significant measures to be included the final logistic model predicting the probability of participating in the assessment. These logistic models are then used to generate predicted probability weights which are included in the estimation of treatment effects. This method ensures that those who participated in the age nine follow-up and have similar characteristics to those who did not are given a larger weight in the analysis.<sup>11</sup>

### C. Data

Unlike previous *PFL* data collection waves which took place in a two and a half year block in order to capture children who were the same age at the time of the interview, data for the age nine follow up were collected in a six-month period between February and June 2019. While the majority of children were nine years old at the time of assessment (average age 9.4 years), the children ranged in age from eight to 11 years old.<sup>12</sup> Importantly, there were no statistically significant differences in the age of the children in the high and low treatment groups at the time of assessment (high treatment=9.5 years, low treatment=9.4 years;  $p$ -value=0.355). As there are no differences in the ages of the children across groups, we do not control for age in the analyses. Data were collected through

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<sup>10</sup> For the child sample, 32 and 11 percent of baseline measures significantly predict attrition from the high and low treatment groups respectively. For the school sample, the corresponding figures are 17 and 19 percent. For the parent sample, the corresponding figures are 36 and 17 percent.

<sup>11</sup> Participants who completed the age nine assessment, but did not complete the baseline assessment, are assigned the average weight.

<sup>12</sup> In particular, 30 percent of the children were eight years old, 42 percent nine years old, 27 percent 10 years old, and 1 percent 11 years old.

direct assessments with children, information provided by schools, and interviews with parents. Detailed information on all instruments used may be found in Appendix B.

### *Direct assessment*

A total of 117 child assessments were conducted lasting approximately 45-60 minutes on average. Children were invited to participate in tests of their cognitive skills using the British Ability Scales III: School Age Battery (BAS III; Elliott, Smith and McCulloch 2011) (an updated version of the assessment used in earlier waves). The BAS III yields an overall score reflecting general cognitive ability (General Conceptual Ability, GCA), as well as three standardized scores for Verbal Ability, Pictorial Reasoning Ability, and Spatial Ability. Children's socio-emotional skills was assessed using the self-reported Social Skills Improvement System Rating Scales (SSIS-RS; Elliot and Gresham, 2008). The majority of these assessments (99%) were conducted in the child's school.

### *School records*

Information on standardized achievement test scores, absenteeism and resource supports were gathered from 32 schools. Data on achievement scores were available for 118 children, while data on absenteeism and school resources were available for 123 children. All schools in Ireland must complete standardized testing for reading and maths in second grade (age seven/eight), and schools can choose to administer the tests more frequently if they wish. The test scores capture children's literacy and numeracy skills based on either the Drumcondra Tests for Reading and Mathematics or the Micra-T for English Reading and Sigma-T for Mathematics tests. Within our sample, 64 percent of children completed the Micra and Sigma tests and 34 percent the Drumcondra tests. The norm-referenced standard scores and categorization of above and below average performance based on STen scores are used as indicators of children's reading and mathematics ability. The absenteeism records include information on school attendance in 2017/18 and 2018/19 based on the number of days in attendance as a proportion of days in the school year. Data was also gathered on the type and amount of additional resource/learning support which the child received.

Table 2 compares the school characteristics of the high and low treatment groups and shows that there are no statistically significant differences across the groups in terms of the proportion of children attending designated disadvantaged status (DEIS) schools<sup>13</sup>, single-gender schools, schools outside the *PFL* catchment area, and the distance of schools from the catchment area. In addition, there are no differences across the high and low treatment groups regarding the proportion of children taking the Drumcondra tests or the Micra/Sigma tests. This is important as it allows us to combine the scores from both tests to estimate treatment effects for the whole sample. Robustness tests, where the

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<sup>13</sup> DEIS schools are classified using socio-economic indicators of the community and school population e.g. unemployment, social housing, literacy levels. These schools receive additional resources including additional staff and funding and access to literacy and numeracy programs.

tests are analyzed separately, are also conducted. Although the *PFL* children attended a wide range of schools (32 in total), 55 percent attended one of the four schools in the *PFL* catchment area. In 26 of the other schools, only one or two *PFL* children were attending. A large proportion of the children were attending a DEIS school (71 percent), 76 percent were attending coeducational schools, and 24 percent were attending single gender schools. The vast majority were attending Catholic schools (98 percent). The finding that the children in the high and low treatment groups are attending similar schools suggests that the program had no impact on parent’s school enrolment decisions.<sup>14</sup>

#### *Parent interview*

In total 111 parent interviews were conducted either online (n=39), over the phone (n=49), or in person (n=23) depending on the participant’s preference. The questionnaire included measures to assess children’s socio-emotional skills using the *Brief Problems Monitor* (BPM; Achenbach *et al.* 2011) and two sub-domains of the *Strengths and Difficulties Questionnaire* (SDQ; Goodman 1997). The interview did not exceed 20 minutes in duration.

#### **D. Estimation of Treatment Effects**

To estimate the effect of *PFL* on age nine outcomes we use permutation-based hypothesis tests with 100,000 replications (see Good 2005; Ludbrook and Dudley 1998; Mewhort 2005).<sup>15</sup> We control for child gender in all specifications as the high treatment group contained more boys than the low treatment group at baseline, and this difference is also present in the age nine samples.<sup>16</sup> We estimate the following equation:

$$Y_i = \beta_0 + \beta_1 D_i + \beta_2 X_i + \epsilon_i$$

where  $Y_i$  is the outcome variable,  $D_i$  is an indicator of treatment status,  $X_i$  represents child gender, and  $\epsilon_i$  is an idiosyncratic error term. The parameter of interest is  $\beta_1$ . As we anticipate the program to have a positive impact on school-age outcomes, one-sided tests with the accepted Type I error rate set at 10 percent are reported. As a robustness test, we add additional covariates to this model to test the impact of controlling for four baseline differences on which the high and low treatment groups differ (i.e. maternal knowledge of child development, parenting self-efficacy, maternal attachment, and maternal consideration of future consequences). When controls are added to the permutation tests, the Freedman–Lane procedure (Freedman and Lane 1983) is used to partition the sample into subsets

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<sup>14</sup> In Ireland, parents are, in theory, free to send their children to a State-funded primary school of their choice. However, due to space constraints, most schools use some form of an admissions policy, generally giving preference to students from the local area. Most children start school between four and five years of age, and all children must be enrolled in school by age six.

<sup>15</sup> As the age nine estimation samples are relatively small, permutation tests which do not depend on the asymptotic behavior of the test statistic are more suitable (Ludbrook and Dudley 1998). This method has been used to estimate treatment effects using previous waves of the *PFL* data (see Doyle *et al.* 2017a; Doyle 2020).

<sup>16</sup> For the age nine child sample, 57 percent of the high treatment group are boys compared to 35 percent of the low treatment group.



which include participants with the same characteristics. Under the null hypothesis of no effect, the outcomes of the high and low treatment groups have the same distributions within a subset. However, when multiple control variables are used, the subsets may become too small, thus each outcome is regressed on the control variables assumed to share a linear relationship with the outcomes. The predicted residuals are then permuted from these regressions within the subsets.

As multiple outcomes are tested at age nine, the stepdown procedure (Romano and Wolf 2005)<sup>17</sup> is used to address issues associated with multiple hypothesis testing. For the cognitive outcomes, the variables are placed into three stepdown categories capturing BAS composite scores, BAS above the norm scores, and BAS below the norm scores. For the school outcomes, variables are placed in four stepdown families capturing standardized test scores, above the norm scores, below the norm scores, and absenteeism/school resources. For the socio-emotional outcomes, the variables are placed into four stepdown categories capturing child-reported socio-emotional scores, child-reported socio-emotional below-average scores, parent-reported socio-emotional scores, and parent-reported socio-emotional below-average scores. The methods used in this paper to estimate treatment effects based on conditional permutation testing with the stepdown procedure, has been used and described in other papers reporting on the *PFL* trial (e.g. Doyle *et al.* 2017a; Doyle 2020).

## IV Results

### A. Cognitive outcomes

The results reported in Table 4 indicate that the *PFL* program continues to have a significant and sizable impact on children's skills at age nine. The treatment increases children's overall BAS GCA score by 0.67 SD, which is similar to the effect of 0.77 SD identified using the same measure at the end of the program (Doyle 2020), demonstrating the sustained impact of the program almost five years after the treatment ended. The results also show that *PFL* has a significant impact on each dimension of cognitive skill including spatial ability (0.48 SD) which involves problem solving, spatial visualization, and short-term visual memory; non-verbal reasoning ability (0.76 SD) which involves inductive reasoning; and also verbal ability (0.39 SD) which involves children's verbal reasoning, verbal knowledge, and expressive language. In addition, all four composite scores survive adjustment for multiple hypotheses testing.

Children are classified as scoring above the norm if their GCA score is greater than 110 points and below the norm if their scores are less than 90 points (range 51-122). Table 4 shows that the high

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<sup>17</sup> First, a *t*-statistic for each null hypothesis in the stepdown family is calculated. Next, the outcome with the largest *t*-statistic is compared with the distribution of maxima permuted *t*-statistics. If the corresponding *p* value is less than 10 percent, the joint null hypothesis is rejected, and the outcome with the lowest *p* value is excluded. The remaining outcomes in the stepdown family are tested again. This process continues until the resulting subset of outcomes fails to be rejected or only one outcome remains.

treatment children are more likely to score *above* the norm in terms of their overall cognitive ability and their spatial ability; however neither result survives multiple hypotheses adjustment with effect sizes ranging from 0.01 to 0.36 SD. In contrast, the high treatment children are less likely to score *below* the norm across all three cognitive domains, as well as overall ability, and the results are robust to multiple hypothesis adjustment. The effect sizes range from 0.42 to 0.67 SD. It is important to note that relatively few children, in either the high or low treatment group, score above the norm, while large proportions score below the norm. For example, only 2 percent of children in the high treatment group have GCA scores above the norm, while 58 percent score below the norm. The BASIII norms are based on a representative UK sample including children across all social groups. The scores identified here thus reflect the disadvantaged nature of the *PFL* cohort, yet the counterfactual reveals that without the *PFL* intervention a significantly greater proportion of the high treatment children would have scored below the norm (78 percent in the low treatment group), thus demonstrating the effectiveness of the program.<sup>18</sup>

The significant results regarding the proportion of children scoring below the norm suggests that the program primarily impacts the lower end of the distribution of children's skills. This is demonstrated in Figure 3 which shows that the distribution of GCA scores for the high treatment group is shifted to the right of the low treatment groups', with larger differences at the lower end of the distribution.

## **B. School outcomes**

Table 5, which reports the results for the school-based outcomes, are in-line with those observed for cognitive skills. The results indicate that the *PFL* program has a significant and large impact on children's academic test scores at age nine, however there are no effects on absenteeism or use of school resource.

The first panel shows the results based on children's second and third class standardized achievement tests. Note that the sample size is lower for third class as complete data are only available for second class as many of the children had not yet entered third class compared. Children in the high treatment group have significantly higher second and third class reading and maths standardized scores to the low treatment group. All the standardized results survive adjustment for multiple hypotheses testing and the effect sizes range from 0.33 to 0.74 SD. Children in the high treatment group are significantly less likely to score *below* the norm in terms of their reading and maths scores in both years, and are significantly more likely to score *above* the norm in terms of their third class reading and maths scores, as well as their second class maths score. Two of the four *above*

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<sup>18</sup> The non-IPW adjusted results are slightly more conservative with somewhat lower effect sizes. For example, 14 of the 16 cognitive outcomes reach conventional levels of significance in the individual IPW-adjusted permutation tests, compared to nine in the non-adjusted results. However, all three stepdown families are still statistically significant in the non-adjusted results.

the norm results survive multiple hypotheses adjustment – third class reading scores and second class maths scores. The effect sizes range from 0.19 to 0.64 SD. All four of the *below* the norm scores survive adjustment for multiple hypothesis testing, with effect sizes of 0.34 to 0.57 SD. In contrast, only one of the six absenteeism and school resources outcomes are statistically significant in the individual tests (high treatment children are less likely to have additional out-of-class educational supports), however it does not survive adjustment for multiple hypothesis testing.<sup>19</sup>

As a robustness test, the school results are re-estimated by conditioning on the schools DEIS (disadvantaged) status and an indicator of whether the child conducted the Drumcondra tests or the Micro/Sigma tests. The findings, provided in Appendix Table B1, show that the results are robust to the inclusion of these controls.<sup>20</sup>

### **C. Child socio-emotional outcomes**

The results in Table 6 indicate that the program has no impact on children’s socio-emotional skills at age nine as measured using the child-reported Social Skills Improvement System Behavior Problems Subscale (SSIS) or the parent-reported Strengths and Difficulties Scale (SDQ) or the Brief Problems Monitor (BPM). Children in the high treatment group have fewer internalizing problems, but more externalizing problems, however these differences are not statistically significant and the effect sizes are low. In addition, although children in the high treatment group have statistically significantly fewer peer problems, attention problems, and total behavioral problems according to the individual permutation tests, these results do not survive adjustment for multiple hypothesis testing. The program also has no impact on the proportion of children with ADHD or learning difficulties. In sum, these results suggest a fade-out of the program’s impact on the socio-emotional and behavioral skills that were observed at age four.<sup>21</sup>

### **D. Conditioning on baseline differences**

As a robustness test, the results are re-estimated conditioning on four variables on which there were significant differences between the high and low treatment groups at baseline and may impact child outcomes – namely maternal knowledge of child development, parenting self-efficacy, maternal attachment, and maternal consideration of future consequences. Appendix Table C2 shows that the

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<sup>19</sup> The non-IPW adjusted results are also similar. There are no differences in the number (or level) of statistically significant results between the IPW and non-IPW adjusted models and the effect sizes are largely similar. However, more of the results survive multiple hypothesis adjustment.

<sup>20</sup> A separate analysis which included all 18 school outcomes in one stepdown family found that the standard reading score in third class, the standard maths score in second class, and the proportion scoring above the norm on maths scores in second class remained statistically significant in the stepdown analysis.

<sup>21</sup> The non-IPW adjusted results are similar to the IPW adjusted results, however two statistically significant differences emerge in the non-IPW adjusted individual permutation tests – children in the high treatment group are significantly less likely to fall within in the above the norm range for bullying problems and overall behavioral problems – however, neither of these results survive multiple hypotheses adjustment.

conditional results for child cognitive outcomes are largely similar to the main results, however two of the outcomes which reach conventional levels of significance in the individual permutation tests are not significant in the conditional results – general conceptual ability and non-verbal above the norm scores. The remaining results are similar in terms of effect size and significance. Table C3 shows that the conditional results for the school outcomes are similar to the main results in terms of the number of statistically significant differences and the effect sizes. Table C4 shows that the conditional results for child socio-emotional outcomes are largely similar to the main results, however one outcome is statistically significant in the conditional results which was not significant in the unconditional results – proportion of children with above the norm behavioral problems – however it does not survive multiple hypothesis adjustment. In addition, the three individual significant treatment effects found in the main results for children’s socio-emotional standardized scores are no longer significant once controls are added. This again confirms a lack of program impact on children’s socio-emotional skills at age nine.

#### **E. Mediation Analysis**

The second aim of this paper is to understand the channels through which these improvements in cognitive skills and achievement were obtained. Thus, a mediation analysis is carried out to examine the potential mechanisms underlying the treatment effects. As discussed in the introduction, *PFL* is a holistic program that aims to impact multiple aspects of parenting in the expectation that these improvements will have a direct impact on child outcomes. Throughout the course of the trial, multiple indicators of parenting were collected between six and 48 months, and 190 of them are considered for inclusion in the mediation analysis.

As the number of potential mediators is greater than the number of participants, 190 OLS models are first conducted to test which variables predict children’s cognitive skills as measured by the GCA standardized score. The models control for gender and are adjusted for IPW as in the main specification. In total, 18 percent (34 of the 190 indicators) are significantly associated with age nine GCA scores. Next, we test which of these 34 measures are significantly impacted by the treatment, again using OLS and controlling for gender and IPW. In the majority of cases, the treatment has no impact on the parenting behaviors that are associated with the GCA scores. However, the treatment does impact the following six variables in the expected direction – the use of electrical socket covers at six months, the HOME (*Home Observation Measurement of the Environment*) Acceptance score at 18 months, the PACOTIS (*Parental Cognition and Conduct Toward the Infant Scale*) Impact Score at 24 months, the PSDQ (*Parenting Styles and Dimensions Questionnaire*) Permissive Parenting score at 36 months, the PSDQ Punitive Parenting score at 36 months, and smoking in the home at 36 months.

In order to summarize these six measures into one aggregate measure of parental investment, exploratory factor analysis (EFA) is applied. The Kaiser criterion, scree plot, (Cattell 1966), and Horn’s (1965) parallel analysis is used to determine the number of factors to be retained. Note that all

six measures have a Kaiser–Meyer–Olkin (KMO) statistic greater than 0.5 and thus are retained (Kaiser 1974). The EFA retained one factor which was reversed scored so that higher values correspond to greater parental investment. Next, a mediation analysis proposed by Imai et al. (2010) is conducted to determine whether, and by how much, the parental investment factor score mediates the relationship between the treatment and children’s age nine GCA scores. Monte Carlo simulation is used to estimate the model parameters for the GCA score and the mediator and confidence intervals are computed based on the average direct effect (ADE) and the average causal mediation effect (ACME) based on simulated data (Hicks and Tingley 2011). All models control for gender and are adjusted for IPW.

The mediation results are reported in Table 7 and include the bootstrap estimates (1000) of the ACME and with 95 percent confidence intervals and the percentage of the effect mediated by the factor. The first column shows that the treatment has a significant effect on the mediator, such that participation in the program raises parental investment as hypothesized. The second column shows that the mediator has a significant impact on children’s GCA scores, such that higher levels of parental investment are associated with better age nine scores. The final model includes both the treatment indicator and the mediator. It shows that the impact of the treatment on children’s scores falls in magnitude once parental investment is controlled for, yet is still statistically significant. In addition, the parental investment mediator is also significant. The bottom panel in the table shows that average causal mediated effect is statistically significant and explains 45.6% of the total effect, thus providing evidence that early parental investment significantly mediates the relationship between the treatment and the GCA score. Thus providing some evidence in support of the *PFL* model.

## V Conclusions

The objective of this study is to determine whether the large and significant impacts of *PFL* found at the end of the program are sustained at age nine. Prior evidence on the medium-term impact of home visiting and parent–training programs in middle childhood is inconclusive, with many studies experiencing a dissolution of effects once the program ends. In contrast, this study finds that *PFL* continues to have a sizeable impact on children’s cognitive skills and achievement tests approximately five years after the families have finished the program. There is no evidence of cognitive fade-out, with effect sizes of 0.67 SD on overall cognitive ability, and significant effects on standardized school achievement tests. In particular, there is an eight-point gap between the high and low treatment groups’ general conceptual ability scores, which is a proxy for IQ. While this is slightly smaller than the 10-point gap found at end of the program, it is still a sizeable difference considering that the low treatment group has received five years of formal schooling.

The program improved all aspects of cognitive skill including spatial ability, non-verbal ability, and verbal ability, in addition to reducing the proportion of children scoring below the

standardized norm. The program is particularly beneficial for children at the lower end of the skills distribution. This contrasts somewhat from earlier findings where there was evidence that the program shifted the entire distribution of cognitive skills. This difference may have occurred as naturally higher ability children, regardless of their treatment status, may be better placed to take advantage of the learning and supports provided in school. The magnitude of the cognitive effects at age nine (0.39 - 0.76 SD) are similar, albeit a little smaller, to those found at the end of the program (0.56 - 0.77 SD). An additional analysis shows that controlling for age five cognitive scores reduces the size of the age nine treatment effects, however the impact of the program on age nine scores is still statistically significant.<sup>22</sup> This suggests that *PFL* is continuing to have an impact on children's development beyond the lifetime of the program.

The results provide evidence in support of the skill formation model (Cunha and Heckman 2007) which posits that developing children's skills early in life helps them to develop more advanced skills later in life through a process of self-productivity, and this in turn raises the effectiveness of later investments, such as investments in schooling. Thus the study offers support for such dynamic complementarities, although an explicit test of this hypothesis is not possible as all children were exposed to schooling. If this process persists and the high treatment group continue to utilize their advanced skill set, this is likely to translate into improved outcomes throughout the life cycle. Indeed, the large and significant treatment effects found for school achievement tests suggest that these advanced cognitive skills are already having an impact on the children's performance in school. Children who received the high treatment supports have better second and third class standardized test scores in reading and maths, with effect sizes ranging from 0.33 - 0.74 SD. The high treatment group is also significantly less likely to score below the norm on their second and third class reading and maths tests and more likely to score above the norm on their third class reading and second class maths tests.<sup>23</sup> Unlike the cognitive tests which were only administered to the *PFL* cohort, the achievement tests were completed by all children in all schools in Ireland as part of national standardized testing, therefore they provide an independent assessment of the children's reading and maths ability. In addition, they allow us to compare the *PFL* cohort to the national norm. In Ireland, about one-third of children score below the norm i.e. a STen score of 4 or below, and one-third of children score above the norm, i.e. a STen score of 7 or above (Department of Education and Skills 2016). In our sample, 31 percent of the high treatment group score *below* the norm on their reading scores and 34 percent on their maths scores, compared to 47 and 55 percent for the low treatment groups' reading and maths scores respectively. In addition, 26 and 22 percent of the high treatment group score *above* the norm on their reading and maths tests, compared to 19 and 5 percent

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<sup>22</sup> Results available upon request.

<sup>23</sup> Note that the third class results are based on a smaller sample size as some of the children had not yet started third class at the time of data collection, thus the third class results should be interpreted with caution.

respectively in the low treatment group. This suggests that the program has raised the ability scores of the high treatment group to the national average, thus eliminating inequalities in school-age skills.

An interesting finding which emerged during data collection was the number and diversity of primary schools attended by the *PFL* cohort. In total, the 123 children are attending 32 different schools. While the majority are attending designated disadvantaged schools (73 and 68 percent in the high and low treatment group respectively), there is still some heterogeneity. Although 55 percent are attending schools within the *PFL* catchment area, many children are attending schools outside the catchment area. As equal proportions of the high and low treatment groups are attending schools outside the catchment area, this suggests that the program did not change parents' preferences regarding the type or location of their child's primary school. From an evaluation perspective it is useful as it suggests that the significant treatment effects found for cognitive ability and achievement tests are unlikely to be driven by differences in school type.

In contrast to the results for children's cognitive skills, the program has no impact on school absenteeism or the use of school resources. The result for absenteeism is not particularly surprising given that most schools have home-school liaison officers who directly address this issue, thus resulting in very low proportions of children missing school (94 percent of the high and low treatment group are present throughout the school year). The rates of absenteeism among the *PFL* cohort is also very similar to the national average rate of absenteeism which ranged from 5.4 to 5.9 percent between 2012/13 and 2016/17 (Millar 2018). The lack of significant treatment effects for the use of additional school resources, such as a special needs assistant and additional literacy and maths support, are more surprising, especially given the high proportion of children scoring below the norm, especially in the low treatment group. Indeed, only 6 and 5 percent of children in the high and low treatment group respectively have a special needs assistant. Although the low treatment group avail of a higher proportion of in-class and out-of-class support (54 and 38 percent respectively) compared to the high treatment group (40 and 24 percent respectively), these differences are not statistically significant. The lack of differences across the two groups may reflect recent policy to provide additional support to all children within the classroom environment rather than singling individual children out for treatment.

The findings also indicate that the significant treatment effects observed for children's socio-emotional skills at age four are no longer present at age nine. At earlier time points we found that the program was effective in reducing the proportion of children within the clinical range of behavioral problems, however, few effects were identified for continuous scores of children's socio-emotional skills (Doyle *et al.* 2016). These earlier measures were based on parent reports only. At age nine we assessed children's socio-emotional skills using both parent and child reports. The lack of treatment effects across both measures, for either the continuous or clinical range, suggests a dissolution of the *PFL*'s earlier impact on children's wellbeing. Given the small effect sizes, these results are unlikely to be driven by sample size issues. The fade-out of these effects may also reflect the smaller number of

significant impacts and lower effect sizes found for non-cognitive skills at earlier time points. These results are in-line with the home visiting literature reported in Appendix Table 1c which finds significant treatment effects in only three of the 10 follow-up studies conducted. The *PFL* treatment effects observed at age four were driven by a small number of children exhibiting significant behavioral problems, thus it is possible that exposure to the school environment, as well as natural child maturation, has helped to reduce these behavioral issues. Indeed, the overall incidence of clinically significant problems at age nine is low in the *PFL* cohort. For example, only 8 percent of the high treatment group and 4 percent of the low treatment group are diagnosed with ASD or ADHD. The higher rate in the high treatment group may reflect the parents' ability to advocate for their children in terms of receiving a diagnosis and/or that the children's conditions were identified by the *PFL* mentors during the program, and thus an early diagnosis was sought. The *PFL* children are now spending a greater proportion of their day outside the home environment, thus it is possible that specific supports targeting the school environment are required to bolster children's development in these areas.

The mediation analysis indicates that ~46 percent of the treatment effect on cognitive skills is explained by improvements in parental investment. While 190 different measures of parental investment were considered as potential mediators, only six were both impacted by the treatment and had an impact on cognitive skills. These measures of investment largely relate to parental style (e.g. engaging in less permissive and punitive parenting) and parental behaviors (e.g. the use of socket covers during infancy and less smoking in the home). Thus the *PFL* model, which is based on providing knowledge and active guidance on appropriate parenting techniques, appears to have been effective, and the cumulative effect of these improvements in parental investment can help to explain the identified effects at age nine. However, to further explore the mechanisms through which the treatment effects emerge, mediation analysis where the treatment effects are decomposed into components associated with parental investments and intervention-induced changes in early skills is needed, as described in Heckman *et al.* (2013) and Heckman *et al.* (2017).

While the effect sizes for the cognitive and achievement tests are larger than those observed in the more contemporary trials, they are in-line with some of the seminal studies that offered home visiting and/or intervened during infancy. For example, our finding of an eight point gap in overall cognitive skills at age nine is similar to the six point gap found in the Jamaica study at age 11/12 (Walker *et al.* 2000). In addition, the standardized effect sizes of 0.67 SD for cognitive skills, 0.33-0.74 SD for reading scores, and 0.47-0.56 SD for maths scores are larger, or of a similar magnitude, to the pooled effect sizes of 0.37 SD for cognition, 0.47 SD for reading, and 0.37 SD for maths found across the age eight, 12, 15, and 21 year follow-ups of the Abecedarian study (Campbell *et al.* 2001). A re-analysis of the Nurse Family Partnership program, which is the most comparable to *PFL* in terms of its prenatal start and its focus on home visiting, by Heckman *et al.* (2017) identify treatment effects



on reading and math tests at age 12, but the effects are restricted to boys and the effect sizes (between 0.15-0.22 SD) are substantially lower than those observed for *PFL*.

These findings raise the question of why the *PFL* treatment effects are larger and more persistent than many other studies in the field, particularly other home visiting programs. The design of the *PFL* program differs in some fundamental ways to the studies mentioned above. In particular it provides five years of multi-treatment supports starting during pregnancy. While some home visiting programs start in the prenatal period (e.g. Nurse Family Partnership), most start after the birth of the child (e.g. Jamaica study). By intervening during pregnancy and the first few months of life, the program capitalizes on heightened brain malleability, but perhaps just as importantly, it allows the *PFL* mentors to build relationships with the treatment families during a unique period of the family's life when they may be more receptive to change. *PFL* is also longer in duration than most home visiting programs. For example, both the Nurse Family Partnership program and the Jamaica study intervene for two-two and a half years, while *PFL* continues until school entry. This is important for two reasons. First, the longer program duration allows the mentors to continue to develop and deepen their relationship with the families, which may result in families becoming more responsive to the treatments provided. Second, building on the skill beget skill hypothesis, intervening from pregnancy until school entry allows the mentors, and consequently the families, to respond appropriately to the child's growing skill set and provide treatment that is tailored to the child's developmental needs. The larger and persistent effects observed in the *PFL* trial may also be attributed to its multi-treatment approach. Although some other programs provide multiple treatments e.g. Abecedarian and Early Head Start offer both home visits and center-based care, *PFL* consistently focused on the parents – teaching parents baby massage, delivering group-based parenting classes, and providing parent-focused home visits. This focus on parents as the primary mechanism of change using different, yet reinforcing, delivery methods may account for the persistence of effects up to age nine.

Although the results reported here increase our understanding of the potential medium-term effects of early and sustained intervention in the 'missing middle' period, the limitations of the study should be noted. In particular, similar to the seminal studies in the field, *PFL* is a single site intervention that took place in one community at one point in time (2008-2015). Thus the findings may not be generalizable to other settings or different populations, however the manualized nature of the program allows for replication and testing in other settings, funding dependant. Another potential issue is spillover effects. As the intervention took place in a small geographical area it is possible that participants in the high and low treatment groups shared *PFL* material or parenting advice, which may result in an under-estimation of the treatment effects (assuming the treatment is effective). However, as tested and reported in earlier studies, a series of questions designed to measure cross talk and information flows across the treatment groups found little evidence of contamination (Doyle *et al.* 2015; Doyle *et al.* 2017, Doyle 2020). Possibly the largest threat to the internal validity concerns attrition. While the retention rate between the age five and age nine follow-ups only fell by about 5

percentage points for the high treatment group, the comparable figure for low treatment groups was about 20 percentage points. Despite this, the groups remained balanced on almost all baseline characteristics, for example, for the child data there are no differences on 109 of the 117 measures considered, thus the types of participants who remained in the sample at age nine are largely equivalent. However, there is some evidence that, across both groups, families with more risk factors were more likely to drop out of the study. To ensure the age nine sample is representative of the original randomized sample, IPW were used in all analysis and as described in the results section, the findings from the IPW are largely similar to the non-IPW results.

To conclude, this study finds that the *PFL* program continues to have a significant impact on important dimensions of children's skills five years after the families finished the program, yet continued investment may be needed to break long-standing inequalities in socio-emotional skills. That said, the sizable cognitive advantages are likely to have positive impacts as the children progress from primary to secondary school, as well as subsequent outcomes in adolescence. Thus it is critical to continue to track the *PFL* cohort as they progress from primary to secondary school. This is particularly important given the magnitude of the cognitive effects, especially in comparison to other intervention programs, as *PFL* can provide a model for other communities aiming to reduce long-term socio-economic inequalities in skills.

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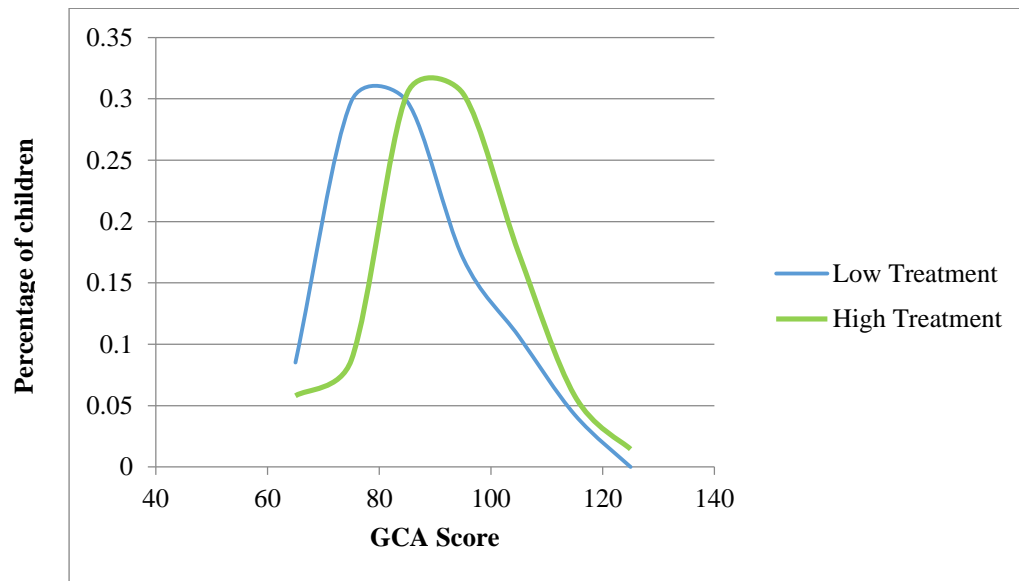


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*Figure 1 Participant flow*



**Figure 2** Distribution of BAS GCA cognitive scores at age 9



**Table 1** Baseline comparison of high and low treatment groups: Child, School, and Parent samples

	Child Sample			School Sample			Parent Sample		
	$M_{\text{HIGH}}$ (SD)	$M_{\text{LOW}}$ (SD)	$p^1$	$M_{\text{HIGH}}$ (SD)	$M_{\text{LOW}}$ (SD)	$p^1$	$M_{\text{HIGH}}$ (SD)	$M_{\text{LOW}}$ (SD)	$p^1$
Age	26.50 (5.59)	25.07 (5.82)	0.197	26.23 (5.67)	25.90 (6.10)	0.763	26.66 (5.69)	26.20 (6.12)	0.694
Married	0.147 (0.36)	0.133 (0.34)	0.890	0.145 (0.36)	0.180 (0.39)	0.515	0.156 (0.37)	0.178 (0.39)	0.782
No. of children	1.97 (1.27)	1.96 (1.07)	0.940	1.91 (1.25)	2.04 (1.07)	0.531	1.98 (1.30)	2.09 (1.10)	0.638
First time mother	0.500 (0.50)	0.422 (0.50)	0.415	0.522 (0.50)	0.380 (0.49)	0.127	0.516 (0.50)	0.378 (0.49)	0.154
Low education (left $\leq$ age 16)	0.294 (0.46)	0.311 (0.47)	0.920	0.304 (0.46)	0.320 (0.47)	0.875	0.328 (0.47)	0.311 (0.47)	0.851
Weschler Abbreviated Scale of Intelligence (WASI)	84.66 (11.35)	81.91 (14.03)	0.275	84.46 (11.38)	82.12 (13.68)	0.328	84.54 (11.55)	83.22 (13.54)	0.600
Employed	0.485 (0.50)	0.444 (0.50)	0.658	0.493 (0.50)	0.440 (0.50)	0.525	0.516 (0.50)	0.467 (0.51)	0.576
Resides in social housing	0.529 (0.50)	0.533 (0.51)	0.877	0.522 (0.50)	0.560 (0.50)	0.666	0.531 (0.50)	0.556 (0.50)	0.777
Medical card	0.574 (0.50)	0.667 (0.48)	0.289	0.565 (0.50)	0.660 (0.48)	0.310	0.547 (0.50)	0.622 (0.49)	0.445
Prior physical health condition	0.794 (0.41)	0.644 (0.48)	0.116	0.783 (0.41)	0.640 (0.49)	<b>0.078</b>	0.781 (0.42)	0.644 (0.48)	0.129
Prior mental health condition	0.309 (0.47)	0.267 (0.45)	0.642	0.290 (0.46)	0.280 (0.45)	0.949	0.281 (0.45)	0.311 (0.47)	0.747
Smoking during pregnancy	0.459 (0.50)	0.422 (0.50)	0.768	0.464 (0.50)	0.440 (0.50)	0.782	0.453 (0.50)	0.467 (0.51)	0.912
Drinking alcohol during pregnancy	0.309 (0.47)	0.311 (0.47)	0.946	0.319 (0.47)	0.300 (0.46)	0.844	0.297 (0.46)	0.333 (0.48)	0.692
<i>N</i>		~113			~119			~109	

**Notes:** All baseline measures were assessed during pregnancy prior to treatment delivery except for WASI which was assessed at 3 months postpartum. Baseline data are missing for four participants in the age 9 assessment. <sup>1</sup> two-tailed  $p$ -values calculated from permutation tests with 100,000 replications.

**Table 2** Comparison of high and low treatment groups: School data

	$M_{\text{HIGH}}$ ( $SD$ )	$M_{\text{LOW}}$ ( $SD$ )	$p^1$
School has DEIS status	0.73 (0.45)	0.68 (0.47)	0.558
Single gender school	0.27 (0.45)	0.21 (0.41)	0.408
School outside <i>PFL</i> catchment area	0.46 (0.50)	0.45 (0.50)	0.976
Driving distance from school to <i>PFL</i> village centre (in kms)	5.51 (14.98)	3.97 (7.25)	0.543
School conducts Micra/Sigma tests	0.67 (0.48)	0.60 (0.50)	0.432
<i>N</i>	70	53	

**Notes:** <sup>1</sup> two-tailed  $p$ -values calculated from permutation tests with 100,000 replications.

**Table 3** Baseline characteristics predicting attrition from the Age 9 Parent sample

	High Treatment Group			Low Treatment Group		
	$M_{\text{STAYER}}$ ( <i>SD</i> )	$M_{\text{NON-STAYER}}$ ( <i>SD</i> )	$p^1$	$M_{\text{STAYER}}$ ( <i>SD</i> )	$M_{\text{NON-STAYER}}$ ( <i>SD</i> )	$p^1$
Age	26.66 (5.69)	23.55 (5.66)	<b>0.006</b>	26.20 (6.12)	24.57 (5.74)	0.179
Married	0.156 (0.37)	0.125 (0.34)	0.679	0.178 (0.39)	0.179 (0.39)	0.897
First time mother	0.516 (0.50)	0.575 (0.50)	0.627	0.378 (0.49)	0.589 (0.50)	<b>0.037</b>
No. of children	1.98 (1.30)	1.88 (1.30)	0.692	2.09 (1.10)	1.77 (1.18)	0.164
Low education (left $\leq$ age 16)	0.328 (0.47)	0.350 (0.48)	0.712	0.311 (0.47)	0.464 (0.50)	0.124
Weschler Abbreviated Scale of Intelligence (WASI)	84.47 (11.32)	78.20 (13.00)	<b>0.013</b>	83.22 (13.54)	79.05 (12.13)	0.114
Employed	0.516 (0.50)	0.125 (0.34)	<b>0.000</b>	0.467 (0.51)	0.339 (0.48)	0.183
Resides in social housing	0.531 (0.50)	0.590 (0.50)	0.569	0.556 (0.50)	0.554 (0.50)	0.845
Medical card	0.547 (0.50)	0.675 (0.47)	0.182	0.622 (0.49)	0.696 (0.46)	0.444
Prior physical health condition	0.781 (0.42)	0.700 (0.46)	0.416	0.644 (0.48)	0.607 (0.49)	0.784
Prior mental health condition	0.281 (0.45)	0.275 (0.452)	0.978	0.311 (0.47)	0.179 (0.39)	0.133
Smoking during pregnancy	0.453 (0.50)	0.600 (0.50)	0.122	0.467 (0.51)	0.482 (0.50)	0.929
Drinking alcohol during pregnancy	0.297 (0.46)	0.175 (0.39)	0.167	0.333 (0.48)	0.214 (0.41)	0.226
<i>N</i>	64	40		45	56	

**Notes:** All baseline measures were assessed during pregnancy prior to treatment delivery except for WASI which was assessed at 3 months postpartum. Baseline data are missing for four participants in the age 9 assessment. <sup>1</sup> two-tailed  $p$ -values calculated from permutation tests with 100,000 replications.

**Table 4** Comparison of high and low treatment groups: Cognitive outcomes

	N (HIGH/LOW)	$M_{HIGH}$ (SD)	$M_{LOW}$ (SD)	$p^1$	$p^2$	ES
<i>BAS Composite Scores</i>						
General Conceptual Ability	116 (69/47)	88.12 (11.85)	80.13 (12.11)	<b>0.002</b>	<b>0.006</b>	0.67
Spatial Ability	117 (69/48)	94.09 (14.26)	86.75 (16.27)	<b>0.032</b>	<b>0.045</b>	0.48
Non-Verbal Ability	117 (69/48)	84.63 (11.67)	76.53 (9.70)	<b>0.000</b>	<b>0.001</b>	0.76
Verbal Ability	116 (69/47)	92.22 (11.70)	87.27 (13.67)	<b>0.043</b>	<b>0.043</b>	0.39
<i>BAS Above the Norm %</i>						
General Conceptual Ability	116 (69/47)	0.02 (0.14)	0.00 (0.00)	<b>0.038</b>	0.239	0.29
Spatial Ability	117 (69/48)	0.11 (0.31)	0.10 (0.31)	0.406	0.679	0.01
Non-Verbal Ability	117 (69/48)	0.03 (0.17)	0.00 (0.00)	<b>0.023</b>	0.156	0.36
Verbal Ability	116 (69/47)	0.05 (0.22)	0.07 (0.26)	0.659	0.659	0.08
<i>BAS Below the Norm %</i>						
General Conceptual Ability	116 (69/47)	0.58 (0.50)	0.78 (0.42)	<b>0.029</b>	<b>0.029</b>	0.42
Spatial Ability	117 (69/48)	0.34 (0.48)	0.66 (0.48)	<b>0.007</b>	<b>0.018</b>	0.67
Non-Verbal Ability	117 (69/48)	0.70 (0.46)	0.89 (0.31)	<b>0.009</b>	<b>0.017</b>	0.51
Verbal Ability	116 (69/47)	0.33 (0.47)	0.60 (0.49)	<b>0.027</b>	<b>0.036</b>	0.58

**Notes:** N' indicates the sample size. 'M' indicates the IPW-adjusted mean. 'SD' indicates the IPW-adjusted standard deviation. <sup>1</sup> one-tailed (right-sided) conditional  $p$ -value from individual IPW-adjusted permutation test with 100,000 replications. <sup>2</sup> one-tailed (right-sided) conditional  $p$ -value from IPW-adjusted stepdown permutation test with 100,000 replications. 'Effect size' is the ratio of the treatment effect to the pooled standard deviation.

**Table 5 Comparison of high and low treatment groups: School outcomes**

	N (HIGH/LOW)	$M_{HIGH}$ (SD)	$M_{LOW}$ (SD)	$p^1$	$p^2$	ES
<i>Academic Standardised Scores</i>						
2 <sup>nd</sup> class reading standardised score	118 (66/52)	99.55 (15.10)	94.93 (12.63)	<b>0.038</b>	<b>0.038</b>	0.33
2 <sup>nd</sup> class maths standardised score	117 (66/51)	97.42 (14.62)	89.73 (12.89)	<b>0.007</b>	<b>0.015</b>	0.56
3 <sup>rd</sup> class reading standardised score	70 (41/29)	97.51 (11.97)	89.67 (9.30)	<b>0.002</b>	<b>0.013</b>	0.74
3 <sup>rd</sup> class maths standardised score	70 (41/29)	94.42 (14.05)	88.05 (12.96)	<b>0.060</b>	<b>0.080</b>	0.47
<i>Academic Above the Norm Cutoff Scores %</i>						
2 <sup>nd</sup> class above average reading score	118 (66/52)	0.26 (0.44)	0.19 (0.39)	0.180	0.180	0.19
2 <sup>nd</sup> class above average maths score	117 (66/51)	0.22 (0.42)	0.05 (0.21)	<b>0.002</b>	<b>0.010</b>	0.55
3 <sup>rd</sup> class above average reading score	70 (41/29)	0.26 (0.44)	0.05 (0.22)	<b>0.006</b>	<b>0.020</b>	0.64
3 <sup>rd</sup> class above average maths score	70 (41/29)	0.14 (0.36)	0.05 (0.22)	<b>0.064</b>	0.164	0.34
<i>Academic Below the Norm Cutoff Scores %</i>						
2 <sup>nd</sup> class below average reading score	118 (66/52)	0.31 (0.47)	0.47 (0.50)	<b>0.043</b>	<b>0.043</b>	0.34
2 <sup>nd</sup> class below average maths score	117 (66/51)	0.34 (0.48)	0.55 (0.50)	<b>0.041</b>	<b>0.087</b>	0.43
3 <sup>rd</sup> class below average reading score	70 (41/29)	0.34 (0.48)	0.62 (0.49)	<b>0.024</b>	<b>0.098</b>	0.57
3 <sup>rd</sup> class below average maths score	70 (41/29)	0.45 (0.50)	0.72 (0.46)	<b>0.042</b>	<b>0.059</b>	0.55
<i>Absenteeism &amp; School Resources %</i>						
Proportion of days present in previous school year	104 (59/45)	0.92 (0.05)	0.93 (0.06)	0.762	0.762	0.20
Proportion of days present in current school year	122 (70/52)	0.94 (0.04)	0.94 (0.06)	0.561	0.763	0.00
In class supports	123 (70/53)	0.40 (0.49)	0.54 (0.50)	0.126	0.325	0.27
Out of class supports	123 (70/53)	0.24 (0.43)	0.38 (0.49)	<b>0.092</b>	0.307	0.31
SNA supports	123 (70/53)	0.06 (0.25)	0.05 (0.21)	0.499	0.798	0.08
Other supports	123 (70/53)	0.11 (0.31)	0.14 (0.35)	0.198	0.606	0.10

**Notes:** N' indicates the sample size. 'M' indicates the IPW-adjusted mean. 'SD' indicates the IPW-adjusted standard deviation. <sup>1</sup> one-tailed (right-sided) conditional  $p$ -value from individual IPW-adjusted permutation test with 100,000 replications. <sup>2</sup> one-tailed (right-sided) conditional  $p$ -value from IPW-adjusted stepdown permutation test with 100,000 replications. 'Effect size' is the ratio of the treatment effect to the pooled standard deviation.



**Table 6** Comparison of high and low treatment groups: Socio-emotional outcomes

	N (HIGH/LOW)	$M_{HIGH}$ (SD)	$M_{LOW}$ (SD)	$p^1$	$p^2$	ES
<i>Child-report socio-emotional skills</i>						
SSIS Internalising Problems	117 (69/48)	7.50 (5.92)	8.24 (5.38)	0.264	0.523	0.13
SSIS Externalising Problems	117 (69/48)	6.42 (4.02)	5.24 (5.60)	0.760	0.760	0.25
SSIS Bullying	117 (69/48)	1.11 (1.23)	0.89 (1.90)	0.606	0.700	0.14
SSIS Hyperactivity/Inattention	117 (69/48)	6.43 (3.53)	6.40 (4.36)	0.428	0.643	0.01
<i>Child-report Socio-emotional skills Cutoff %</i>						
SSIS Internalising Problems	117 (69/48)	0.10 (0.31)	0.10 (0.30)	0.505	0.505	0.02
SSIS Externalising Problems	117 (69/48)	0.07 (0.26)	0.14 (0.35)	0.285	0.377	0.22
SSIS Bullying	117 (69/48)	0.01 (0.10)	0.10 (0.31)	0.112	0.177	0.45
SSIS Hyperactivity/Inattention	117 (69/48)	0.12 (0.33)	0.11 (0.32)	0.495	0.625	0.01
<i>Parent-report socio-emotional skills</i>						
SDQ Peer problems	111 (64/47)	1.59 (1.61)	2.11 (1.73)	<b>0.060</b>	0.209	0.31
SDQ Prosocial behaviour	111 (64/47)	8.85 (1.71)	8.78 (1.34)	0.279	0.490	0.05
BPM Internalising problems	110 (63/47)	57.61 (6.35)	57.74 (6.13)	0.348	0.348	0.02
BPM Externalising problems	110 (63/47)	53.79 (5.74)	54.79 (6.25)	0.282	0.437	0.17
BPM Attention problems	110 (63/47)	53.89 (7.46)	56.71 (7.10)	<b>0.038</b>	0.136	0.39
<i>Parent-report socio-emotional cutoff scores %</i>						
SDQ Peer problems	111 (64/47)	0.13 (0.34)	0.18 (0.39)	0.224	0.508	0.15
SDQ Prosocial behaviour	111 (64/47)	0.08 (0.27)	0.01 (0.12)	0.899	0.899	0.34
BPM Internalising problems	110 (63/47)	0.21 (0.41)	0.24 (0.43)	0.244	0.566	0.07
BPM Externalising problems	110 (63/47)	0.07 (0.27)	0.18 (0.39)	0.187	0.325	0.31
BPM Attention problems	110 (63/47)	0.13 (0.33)	0.15 (0.36)	0.350	0.459	0.07
<i>Other</i>						
SISS Problem Behaviours Standardised Total Score	117 (69/48)	96.65 (10.20)	97.42 (14.64)	0.381	~	0.06
SISS Problem Behaviours ~ 85%ile	117 (69/48)	0.04 (0.20)	0.10 (0.30)	0.869	~	0.23
BPM Total problems standardised score	110 (63/47)	54.60 (6.87)	56.70 (6.46)	<b>0.079</b>	~	0.32
Child has ASD-ADHD %	111 (64/47)	0.08 (0.27)	0.04 (0.20)	0.622	~	0.14
Child has learning difficulty %	111 (64/47)	0.06 (0.23)	0.12 (0.33)	0.177	~	0.22

**Notes:** N' indicates the sample size. 'M' indicates the IPW-adjusted mean. 'SD' indicates the IPW-adjusted standard deviation. <sup>1</sup> one-tailed (right-sided) conditional  $p$ -value from individual IPW-adjusted permutation test with 100,000 replications. <sup>2</sup> one-tailed (right-sided) conditional  $p$ -value from IPW-adjusted stepdown permutation test with 100,000 replications. 'Effect size' is the ratio of the treatment effect to the pooled standard deviation.

**Table 7** Cognitive skills results – Mediation analysis

	Parental Investment	GCA Score	GCA Score
Treatment	1.113*** (0.275)	9.642*** (3.111)	5.201* (2.950)
Parental investment	~	~	3.990*** (1.199)
ACME			4.44 [1.48, 8.51]
Direct effect			5.33 [-.01, 10.90]
Total effect			9.77 [4.17, 15.67]
% of total effect mediated			45.6%

**Note: N=95.** All models include child gender and are adjusted for attrition using IPW weights. Rounded brackets show robust standard errors. Square brackets show 95 percent confidence intervals. \*\*\*p < 0.01, \*\*p < 0.05, \*p < 0.10.

## Appendix A

**Table 1a** Evaluations of Cognitive Outcomes for Home Visiting Programs from Ages 5-12

Outcome	Author	Sample Size	Programme	Measures	Significant Finding	Effect	Age (years)	% of original sample retained
Cognitive tests	Bierman <i>et al.</i> (2017)	556 children	Early Head Start Home Based Option + REDI-P	Woodcock Johnson tests of Achievement (Letter-Word Identification Scale), School Readiness Questionnaire, Learning Behaviours Scale	Woodcock Johnson tests of Achievement	Favorable	7-9	87%
	Chazan-Cohen, Raikes and Vogel (2013)	927 families	Early Head Start Home Based Option	English receptive vocabulary, Woodcock-Johnson revised test, Leiter R sustained attention test, speech problems survey	None	None	5	Original N not reported. Retention between ages 2-5: 96% parent interviews, 92% child assessments
	Kitzman <i>et al.</i> (2010)	635 children	Nurse Family Partnership	Leiter-R sustained attention test	None	None	12	80% parent interviews, 76% child interviews, 85% school records
	Olds <i>et al.</i> (2014)	411 children	Nurse Family Partnership	Conners continuous performance test (attention dysfunction)	None	None	6-9	81% at age 6 78% at age 9

**Table 1b** *Evaluations of Achievement Test Outcomes for Home Visiting Programs from Ages 5-12*

Outcome	Author	Sample Size	Programme	Measures	Significant Finding	Effect	Age (years)	% of original sample retained
Achievement tests	Bierman <i>et al.</i> (2017)	556 children	Early Head Start Home Based Option + REDI-P	Test of Word Rereading Efficiency, Academic Competence Evaluation Scales (ACES; reading/language skills, mathematics), Academic Performance Rating Scale	ACES (reading/language skills)	Favorable	7-9	87%
	DuMont <i>et al.</i> (2010)	897 mothers	Healthy Families America	Partaking in a gifted program, receiving remedial services, receiving special education, repeating a grade	Participating in a gifted programme, receiving special education	Favorable	7	80% of baseline sample
	Kirkland and Mitchell-Herzfield (2012)	577 mother and child pairs	Healthy Families America	Doing poorly academically (below grade level in reading, math or positive behaviors that promote learning), excelling academically (reading and math, behaviors that promote learning), retained in 1 <sup>st</sup> grade	Excelling academically with all 3 behaviors that promote learning, retained in 1 <sup>st</sup> grade	Favorable	6-7	49% academic reports, 68% child interviews, 80% parent surveys
	Kitzman <i>et al.</i> (2010)	635 children	Nurse Family Partnership	Placement in special education, ever retained in a grade, GPA, Group achievement test scores, Peabody Individual Achievement Tests	None	None	12	80% parent interviews, 76% child interviews, 85% school records
	Sidora-Arcoleo <i>et al.</i> (2010)	721 mother and child dyads	Nurse Family Partnership	Peabody Picture Vocabulary Test-Revised	None	None	6-12 years	Not reported

**Table 1c** Evaluations of Socio-emotional Outcomes for Home Visiting Programs from Ages 5-12

Outcome	Author	Sample Size	Programme	Measures	Significant Finding	Effect	Age (years)	% of original sample retained
<b>Socio-emotional development</b>	Bierman <i>et al.</i> (2017)	556 children	Early Head Start Home Based Option + REDI-P	Social Competence Scale, Student Teacher Relationship Scale, Child Behavior Scale (excluded by peers), Perceived Competence Scale for Children, Loneliness Scale, Friendship Questionnaire	Perceived Competence Scale for Children, Child Behaviour Scale	Favorable	7-9	87%
	Chazan-Cohen, Raikes and Vogel (2013)	927 families	Early Head Start Home Based Option	Child Behavior Checklist (CBCL), Family and Child Experiences Survey (FACES), negativity towards parents during play	FACES (positive approaches to learning), FACES (social behavior problems)	Favorable	5	Original N not reported. Retention between ages 2-5: 96% parent interviews 92% child assessments
	Dumont <i>et al.</i> (2010)	897 mothers	Healthy Families America	CBCL (attention problems, rule breaking and aggressive behaviours, social problems, and the anxious-depressed and withdrawn-depressed syndrome)	None	None	7	80% of baseline sample
	Kitzman <i>et al.</i> (2010)	635 children	Nurse Family Partnership	Child conduct	None	None	12	80% parent interviews, 76% child interviews, 85% school records
	Kitzman <i>et al.</i> , (2010)	594 mothers, 578 children	Nurse Family Partnership	Ever arrested, externalizing disorders and internalizing disorders, total problems	Internalizing disorders	Favorable	12	80% parent interviews, 76% child interviews, 85% school records
	Minkowitz <i>et al.</i> (2007)	1308 children	Healthy Steps	CBCL (internalizing, externalizing, total problems)	None	None	5.5	57%
	Olds <i>et al.</i> (2014)	411 children	Nurse Family Partnership	CBCL (internalizing, externalizing, total problems)	None	None	6-9	81% at age 6, 78% at age 9

	Sidora-Arcoleo <i>et al.</i> (2010)	721 mother and child dyads	Nurse Family Partnership	CBCL (physical aggression items)	None	None	6-12	Not reported
	Sitnick <i>et al.</i> (2015)	614 families	Family Check Up for Children	CBCL (oppositional-aggressive items)	None	None	5	85%
	Smith <i>et al.</i> (2014)	612 children	Family Check Up for Children	CBCL (oppositional-aggressive items), oppositional behavior in the classroom	None	None	7-8	62%

## Appendix B Measure Description

### Cognitive skills

Children's cognitive ability was assessed using the School Age *British Ability Scales III* (Elliot and Smith 2011). The BAS III School Age battery was designed as an assessment of children's abilities in clinical, educational, and research settings for children and students aged from 5 years 0 months to 17 years 11 months. The BAS III consists of six subscales: word definitions, verbal similarities, matrices, quantitative reasoning, recognition of designs, and pattern construction. These sub-scales yield an overall score reflecting general cognitive ability (General Conceptual Ability, GCA), as well as three cluster scores for Verbal Ability, Non-Verbal Ability, and Spatial Ability. The GCA score assesses overall cognitive ability such as thinking logically, making decisions, and learning. The Spatial Ability score assesses problem solving, spatial visualisation, and short-term visual memory. The Nonverbal Reasoning score assesses inductive reasoning. The Verbal ability score assesses children's verbal reasoning, verbal knowledge, and expressive language. Age-based *T* scores are calculated for each domain that are standardised to have a mean of 100 and a standard deviation of 15, as well as cutoff scores indicating whether the child scores above or below average for the GCA and cluster scores.

### Socio-emotional skills

The *Social Skills Improvement System Rating Scales* (SSIS-RS; Elliot and Gresham 2008) is a child report measure ( $\alpha = 0.87$ ) that assesses children's social skills and problem behaviour with two subscales by those names. The current study used the behaviour problems subscale, which consists of 29-items measuring: externalising (x items;  $\alpha = 0.79$ ), internalising (x items;  $\alpha = 0.84$ ), bullying (x items;  $\alpha = 0.44$ ), and hyperactivity/inattention (x items;  $\alpha = 0.70$ ) for children aged 3 to 18 years. Children indicated how true a statement about each social skill and problem behaviour was for them using a 4-point scale of *not true*, *a little true*, *a lot true*, and *very true*. These were scored as 0, 1, 2, 3 respectively. The relevant items were summed to create subscale scores. A total score was also computed and converted to a standard score using the scale norms for age and gender. Cutoff scores were also created for each of the subscales and the standardised total score to indicate whether or not the child scored above average for each score.

### Achievement Test Scores

All primary schools in Ireland are required to administer standardised achievement tests in English reading and Mathematics at the end of second class when children are approximately 7/8 years of age, some school also continue to administer standardised achievement tests at the end of each school class. There are two published sets of standardised achievement tests normed for the Irish population that primary schools can use: Drumcondra Tests for Reading and Mathematics, produced by the Educational Research Centre or the Micra-T for English Reading and Sigma-T for Mathematics produced by Folens. These norm-referenced tests provide a standard score (mean = 100, SD = 15), with higher scores indicative of better performance. Standard scores are also converted to STen scores ranging from 1 to 10 indicating the child's approximate position with respect to the reference population. Standard and STen scores for participating children were obtained from the child's school. Scores for 2<sup>nd</sup> and 3<sup>rd</sup> class, where available (equivalent to US 2<sup>nd</sup> and 3<sup>rd</sup> grade and UK Year 3 and 4) were used in this report. STen scores were converted to indicators of above and below average performance based on National Council for Curriculum and Assessment (NCCA) STen score descriptors.

### Absenteeism and School Resources

Information on *daily school attendance* was gathered from participating children's schools in the form of days attended and days absent for the previous school year (2017/2018) and up to the point of data collection for the year of data collection (2018/2019). This information was used to calculate the proportion of days attended in the school year for the current and previous school year.

Children with special education needs in Primary Schools in Ireland receive educational support to assist their integration into ordinary mainstream schools or to support their education in specialised schools. These supports normally take the form of in/out of class resource/learning support, or Special Needs Assistant depending on the child's assessed level of need. Schools were asked to indicate, yes or no, whether each child received education supports in school in the form of in-class support, out of class support, Special Needs Assistant, or other educational supports.

### *Parent-reported Instruments*

#### Socio-emotional Skills

The *Brief Problems Monitor* (BPM; Achenbach, McConaughy, Ivanovaa, and Rescorla 2011) is a parent report measure for children aged 6-18 years to monitor children's functioning and responses to interventions. The BPM is based on items from the Child Behaviour Checklist (CBCL), Teachers Report Form (TRF) and Youth Self Report (YSR). The measure consists of 19 items with the response options *not true*, *somewhat true*, or *very true*. These were scored as zero, 1 and 2 respectively. The measure yields scores across three subscales: internalising (6 items;  $\alpha = 0.76$ ), attention (6 items;  $\alpha = 0.89$ ) and externalising (7 items;  $\alpha = 0.80$ ) problems. The scores for each of the three problems subscales were summed to create a Total Problems score. Scores were then converted to standard scores based on the child's age and gender, and binary indicators of concerning problem behaviour were created based on standard scores exceeding 60.

The *Strengths and Difficulties Questionnaire* (SDQ; Goodman 1997) is a 25-item questionnaire assessing behaviours, emotions, and relationships of 4- to 16-year-olds. The questionnaire covers five dimensions: conduct problems, emotional symptoms, hyperactivity, peer problems, and pro-social behaviour. The 5-item Peer Problems ( $\alpha = 0.62$ ) and 5-item Pro-Social ( $\alpha = 0.72$ ) subscales were used in this report. Items were scored 0 for *not true*, 1 for *somewhat true*, and 2 for *certainly true*. Two items from the Peer Problems subscale were reverse scored. The five items for each subscale were summed giving a total score of 0 to 10 for each subscale ( $\alpha = 0.90$ ). Cutoff scores were also created to indicate scores that were of clinical concern.

Binary indicators for whether or not the child has a diagnosis of Autism Spectrum Disorder/Attention Deficit Hyperactivity Disorder (*ASD-ADHD*) or whether or not the child has a diagnosed *learning difficulty* were created from parents responses to a question asking them if their child has a developmental delay/disorder or any diagnosed learning or physical disability, and if so to list the diagnosis.



## Appendix C Robustness Tests

**Table C1** Comparison of high and low treatment groups: School outcomes conditional on DEIS status and type of test

	N (HIGH/LOW)	$M_{\text{HIGH}}$ (SD)	$M_{\text{LOW}}$ (SD)	$p^1$	$p^2$	ES
<i>Academic Standardised Scores</i>						
2 <sup>nd</sup> class reading standardised score	118 (66/52)	99.55 (15.10)	94.93 (12.63)	<b>0.044</b>	<b>0.044</b>	0.33
3 <sup>rd</sup> class reading standardised score	69 (40/29)	97.51 (12.09)	89.67 (9.30)	<b>0.001</b>	<b>0.011</b>	0.73
2 <sup>nd</sup> class maths standardised score	117 (66/51)	97.42 (14.62)	89.73 (12.89)	<b>0.003</b>	<b>0.011</b>	0.56
3 <sup>rd</sup> class maths standardised score	69 (40/29)	94.47 (14.19)	88.05 (12.96)	<b>0.058</b>	<b>0.081</b>	0.47
<i>Academic Above the Norm Cutoff Scores</i>						
2 <sup>nd</sup> class above average reading score %	118 (66/52)	0.26 (0.44)	0.19 (0.39)	0.127	0.127	0.19
3 <sup>rd</sup> class above average reading score %	69 (40/29)	0.26 (0.44)	0.05 (0.22)	<b>0.006</b>	<b>0.014</b>	0.65
2 <sup>nd</sup> class above average maths score %	117 (66/51)	0.22 (0.42)	0.05 (0.21)	<b>0.000</b>	<b>0.006</b>	0.55
3 <sup>rd</sup> class above average maths score %	69 (40/29)	0.15 (0.36)	0.05 (0.22)	<b>0.059</b>	0.119	0.35
<i>Academic Below the Norm Cutoff Scores</i>						
2 <sup>nd</sup> class below average reading score %	118 (66/52)	0.31 (0.47)	0.47 (0.50)	<b>0.044</b>	<b>0.044</b>	0.34
3 <sup>rd</sup> class below average reading score %	69 (40/29)	0.35 (0.48)	0.62 (0.49)	<b>0.033</b>	<b>0.047</b>	0.56
2 <sup>nd</sup> class below average maths score %	117 (66/51)	0.34 (0.48)	0.55 (0.50)	<b>0.030</b>	<b>0.075</b>	0.43
3 <sup>rd</sup> class below average maths score %	69 (40/29)	0.44 (0.50)	0.72 (0.46)	<b>0.035</b>	<b>0.091</b>	0.57
<i>Absenteeism &amp; School Resources</i>						
Proportion of days present in last school year %	100 (55/45)	0.93 (0.05)	0.93 (0.06)	0.663	0.663	0.11
Proportion of days present in current school year %	117 (66/51)	0.94 (0.04)	0.94 (0.06)	0.626	0.746	0.04
In class supports %	118 (66/52)	0.39 (0.49)	0.53 (0.50)	<b>0.092</b>	0.400	0.28
Out of class supports %	118 (66/52)	0.24 (0.43)	0.37 (0.49)	0.121	0.359	0.28
SNA supports %	118 (66/52)	0.04 (0.19)	0.03 (0.17)	0.279	0.709	0.04
Other supports %	118 (66/52)	0.09 (0.29)	0.14 (0.35)	0.158	0.556	0.16

**Notes:** N' indicates the sample size. 'M' indicates the IPW-adjusted mean. 'SD' indicates the IPW-adjusted standard deviation. <sup>1</sup> one-tailed (right-sided) conditional  $p$ -value from individual IPW-adjusted permutation test with 100,000 replications. <sup>2</sup> one-tailed (right-sided) conditional  $p$ -value from IPW-adjusted stepdown permutation test with 100,000 replications. 'Effect size' is the ratio of the treatment effect to the standard deviation of the low treatment group. The conditional set includes gender, DEIS status, and type of test conducted.

**Table C2 Comparison of high and low treatment groups: Child cognitive outcomes conditional on baseline differences**

	N (HIGH/LOW)	$M_{HIGH}$ (SD)	$M_{LOW}$ (SD)	$p^1$	$p^2$	ES
<i>BAS Composite Scores</i>						
General Conceptual Ability	110 (68/42)	88.18 (11.95)	79.65 (12.85)	<b>0.008</b>	<b>0.022</b>	0.61
Spatial Ability	111 (68/43)	94.39 (14.22)	86.72 (17.08)	<b>0.059</b>	0.129	0.42
Non-Verbal Ability	111 (68/43)	84.53 (11.76)	76.47 (10.22)	<b>0.000</b>	<b>0.004</b>	0.73
Verbal Ability	110 (68/42)	92.19 (11.81)	86.19 (13.79)	<b>0.054</b>	0.114	0.39
<i>BAS Above the Norm %</i>						
General Conceptual Ability	110 (68/42)	0.02 (0.14)	0.00 (0.00)	0.190	0.473	0.31
Spatial Ability	111 (68/43)	0.11 (0.31)	0.12 (0.33)	0.521	0.789	0.03
Non-Verbal Ability	111 (68/43)	0.03 (0.18)	0.00 (0.00)	0.241	0.460	0.25
Verbal Ability	110 (68/42)	0.05 (0.22)	0.08 (0.28)	0.845	0.845	0.19
<i>BAS Below the Norm %</i>						
General Conceptual Ability	110 (68/42)	0.57 (0.50)	0.74 (0.44)	<b>0.041</b>	<b>0.041</b>	0.38
Spatial Ability	111 (68/43)	0.33 (0.47)	0.68 (0.47)	<b>0.008</b>	<b>0.015</b>	0.68
Non-Verbal Ability	111 (68/43)	0.71 (0.46)	0.88 (0.33)	<b>0.006</b>	<b>0.023</b>	0.52
Verbal Ability	110 (68/42)	0.33 (0.47)	0.66 (0.48)	<b>0.024</b>	<b>0.030</b>	0.54
<i>NIH Toolbox Executive Functioning</i>						
Flanker Task - Inhibitory control	110 (68/42)	97.89 (16.77)	90.11 (11.90)	<b>0.061</b>	<b>0.061</b>	0.39
Dimensional Change Card Sort Task - Attention flexibility	109 (68/41)	102.26 (21.88)	91.76 (12.57)	<b>0.046</b>	<b>0.065</b>	0.44
List Sorting Task - Working memory	107 (68/39)	96.05 (13.46)	89.51 (9.46)	<b>0.011</b>	<b>0.069</b>	0.49
<i>Other</i>						
Composite Executive Function score	107 (68/39)	0.38 (1.03)	0.18 (0.69)	<b>0.033</b>	<b>0.033</b>	0.49

**Notes:** N' indicates the sample size. 'M' indicates the IPW-adjusted mean. 'SD' indicates the IPW-adjusted standard deviation. <sup>1</sup> one-tailed (right-sided) conditional  $p$ -value from individual IPW-adjusted permutation test with 100,000 replications. <sup>2</sup> one-tailed (right-sided) conditional  $p$ -value from IPW-adjusted stepdown permutation test with 100,000 replications. 'Effect size' is the ratio of the treatment effect to the pooled standard deviation. The conditional set includes gender maternal knowledge of child development, parenting self-efficacy, maternal attachment, and maternal consideration of future consequences.

**Table C3 Comparison of high and low treatment groups: School outcomes conditional on baseline differences**

	N (HIGH/LOW)	$M_{HIGH}$ (SD)	$M_{LOW}$ (SD)	$p^1$	$p^2$	ES
<i>Academic Standardised Scores</i>						
2 <sup>nd</sup> class reading standardised score	112 (65/47)	99.89 (15.06)	95.10 (13.08)	<b>0.061</b>	<b>0.061</b>	0.32
3 <sup>rd</sup> class reading standardised score	66 (41/25)	97.51 (11.97)	89.13 (9.58)	<b>0.009</b>	<b>0.028</b>	0.65
2 <sup>nd</sup> class maths standardised score	111 (65/46)	97.49 (14.77)	89.99 (13.36)	<b>0.011</b>	<b>0.026</b>	0.50
3 <sup>rd</sup> class maths standardised score	66 (41/25)	94.42 (14.05)	86.22 (12.98)	<b>0.037</b>	<b>0.054</b>	0.53
<i>Academic Above the Norm Cutoff Scores</i>						
2 <sup>nd</sup> class above average reading score %	112 (65/47)	0.27 (0.45)	0.21 (0.41)	0.283	0.283	0.12
3 <sup>rd</sup> class above average reading score %	66 (41/25)	0.26 (0.44)	0.05 (0.23)	<b>0.014</b>	<b>0.052</b>	0.52
2 <sup>nd</sup> class above average maths score %	111 (65/46)	0.22 (0.42)	0.05 (0.22)	<b>0.005</b>	<b>0.026</b>	0.49
3 <sup>rd</sup> class above average maths score %	66 (41/25)	0.14 (0.36)	0.05 (0.23)	0.128	0.277	0.27
<i>Academic Below the Norm Cutoff Scores</i>						
2 <sup>nd</sup> class below average reading score %	112 (65/47)	0.30 (0.46)	0.46 (0.50)	<b>0.067</b>	<b>0.067</b>	0.30
3 <sup>rd</sup> class below average reading score %	66 (41/25)	0.34 (0.48)	0.63 (0.49)	<b>0.020</b>	<b>0.049</b>	0.61
2 <sup>nd</sup> class below average maths score %	111 (65/46)	0.35 (0.48)	0.52 (0.50)	<b>0.063</b>	<b>0.099</b>	0.34
3 <sup>rd</sup> class below average maths score %	66 (41/25)	0.45 (0.50)	0.80 (0.41)	<b>0.015</b>	<b>0.044</b>	0.65
<i>Absenteeism &amp; School Resources</i>						
Proportion of days present in last school year %	99 (59/05)	0.92 (0.05)	0.94 (0.06)	0.708	0.708	0.15
Proportion of days present in current school year %	116 (69/47)	0.94 (0.04)	0.94 (0.06)	0.401	0.579	0.10
In class supports %	117 (69/48)	0.39 (0.49)	0.56 (0.50)	<b>0.090</b>	0.276	0.31
Out of class supports %	117 (69/48)	0.24 (0.43)	0.40 (0.49)	0.076	0.273	0.34
SNA supports %	117 (69/48)	0.07 (0.25)	0.03 (0.18)	0.618	0.871	0.13
Other supports %	117 (69/48)	0.09 (0.29)	0.13 (0.34)	0.237	0.625	0.11

**Notes:** N' indicates the sample size. 'M' indicates the IPW-adjusted mean. 'SD' indicates the IPW-adjusted standard deviation. <sup>1</sup> one-tailed (right-sided) conditional  $p$ -value from individual IPW-adjusted permutation test with 100,000 replications. <sup>2</sup> one-tailed (right-sided) conditional  $p$ -value from IPW-adjusted stepdown permutation test with 100,000 replications. 'Effect size' is the ratio of the treatment effect to the standard deviation of the low treatment group. The conditional set includes gender maternal knowledge of child development, parenting self-efficacy, maternal attachment, and maternal consideration of future consequences.

**Table C4 Comparison of high and low treatment groups: Child socio-emotional outcomes conditional on baseline differences**

	N (HIGH/LOW)	$M_{HIGH}$ (SD)	$M_{LOW}$ (SD)	$p^1$	$p^2$	ES
<i>Child-report</i>						
SSIS Internalizing Problems	111 (68/43)	7.54 (5.83)	8.36 (5.58)	0.304	0.552	0.09
SSIS Externalizing Problems	111 (68/43)	6.43 (3.92)	4.89 (5.61)	0.830	0.887	0.25
SSIS Bullying	111 (68/43)	1.11 (1.24)	0.60 (1.68)	0.818	0.818	0.27
SSIS Hyperactivity/Inattention	111 (68/43)	6.45 (3.48)	6.17 (4.40)	0.585	0.789	0.08
<i>Child-report Cutoff %</i>						
SSIS Internalizing Problems	111 (68/43)	0.11 (0.31)	0.11 (0.32)	0.376	0.376	0.04
SSIS Externalizing Problems	111 (68/43)	0.08 (0.27)	0.09 (0.29)	0.467	0.559	0.04
SSIS Bullying	111 (68/43)	0.01 (0.10)	0.04 (0.21)	0.178	0.456	0.17
SSIS Hyperactivity/Inattention	111 (68/43)	0.11 (0.31)	0.13 (0.34)	0.346	0.521	0.09
<i>Parent-report</i>						
SDQ Peer problems	107 (64/43)	1.59 (1.61)	1.93 (1.49)	0.115	0.385	0.20
SDQ Prosocial behaviour	107 (64/43)	8.85 (1.71)	8.80 (1.29)	0.465	0.456	0.05
BPM Internalising problems	106 (63/43)	57.54 (6.34)	57.23 (6.14)	0.452	0.663	0.04
BPM Externalising problems	106 (63/43)	53.79 (5.74)	54.70 (6.08)	0.261	0.398	0.17
BPM Attention problems	106 (63/43)	53.89 (7.46)	56.47 (7.05)	0.133	0.370	0.23
<i>Parent-report Cutoff%</i>						
SDQ Peer problems	107 (64/43)	0.13 (0.34)	0.14 (0.35)	0.384	0.762	0.03
SDQ Prosocial behaviour	107 (64/43)	0.08 (0.27)	0.02 (0.12)	0.918	0.918	0.37
BPM Internalising problems	106 (63/43)	0.21 (0.41)	0.22 (0.42)	0.344	0.695	0.02
BPM Externalising problems	106 (63/43)	0.07 (0.27)	0.16 (0.37)	0.262	0.491	0.23
BPM Attention problems	106 (63/43)	0.13 (0.33)	0.14 (0.35)	0.531	0.657	0.03
<i>Other</i>						
SISS Problem Behaviours Standardised Total Score	111 (68/43)	96.73 (10.03)	96.63 (14.85)	0.506	0.506	0.02
SISS Problem behaviours total score - ~ 85%ile	111 (68/43)	0.04 (0.20)	0.12 (0.32)	<b>0.072</b>	0.205	0.32
BPM Total problems standardised score	106 (63/43)	54.60 (6.87)	56.48 (6.44) <sup>3</sup>	0.138	0.256	0.24
Child has ASD-ADHD	107 (64/43)	0.08 (0.27)	0.05 (0.21)	0.561	0.561	0.10
Child has learning difficulty	107 (64/43)	0.06 (0.23)	0.11 (0.32)	0.204	0.260	0.19

**Notes:** N' indicates the sample size. 'M' indicates the IPW-adjusted mean. 'SD' indicates the IPW-adjusted standard deviation. <sup>1</sup> one-tailed (right-sided) conditional  $p$ -value from individual IPW-adjusted permutation test with 100,000 replications. <sup>2</sup> one-tailed (right-sided) conditional  $p$ -value from IPW-adjusted stepdown permutation test with 100,000 replications. 'Effect size' is the ratio of the treatment effect to the pooled standard deviation. The conditional set includes gender maternal knowledge of child development, parenting self-efficacy, maternal attachment, and maternal consideration of future consequences.

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